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# Proceedings of International Workshop: Constructive Mathematical Analysis

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KONYA, TÜRKİYE

**CMA**  
CONSTRUCTIVE MATHEMATICAL ANALYSIS

## Message from Editor

Dear Participants,

It's my pleasure to edit the Proceedings Book of "2nd International Workshop: Constructive Mathematical Analysis". Our workshop is an activity of the journal Constructive Mathematical Analysis and supported by Scientific Research Projects Coordinatorship of Selçuk University. By organizing this workshop, our main aim was to promote, encourage, and provide a forum for the academic exchange of ideas and recent research works. The workshop presents new results and future challenges, in a series of virtual keynote lectures and virtual contributed short talks. In our workshop, we provide a forum for mathematicians to communicate recent research results in the areas of Real analysis, Complex analysis, Potential theory, Special functions and their applications, Matrix analysis, Approximations and expansions, Harmonic analysis, Fourier analysis, Integral transforms-operational calculus, Functional analysis, Fixed Point Theory, Operator theory, Miscellaneous applications of functional analysis, Convex and geometric analysis, Stochastic analysis, Numerical analysis, and the Applications of these fields in other areas. The conference was face to face, and only one presentation was allowed for each participant. The language of all presentations was English, and submissions were peer-reviewed by at least two referees. This proceedings book also includes the papers which are peer-reviewed by at least two referees.

Thanks.

**Assoc. Prof. Dr. Tuncer Acar**  
Selçuk University  
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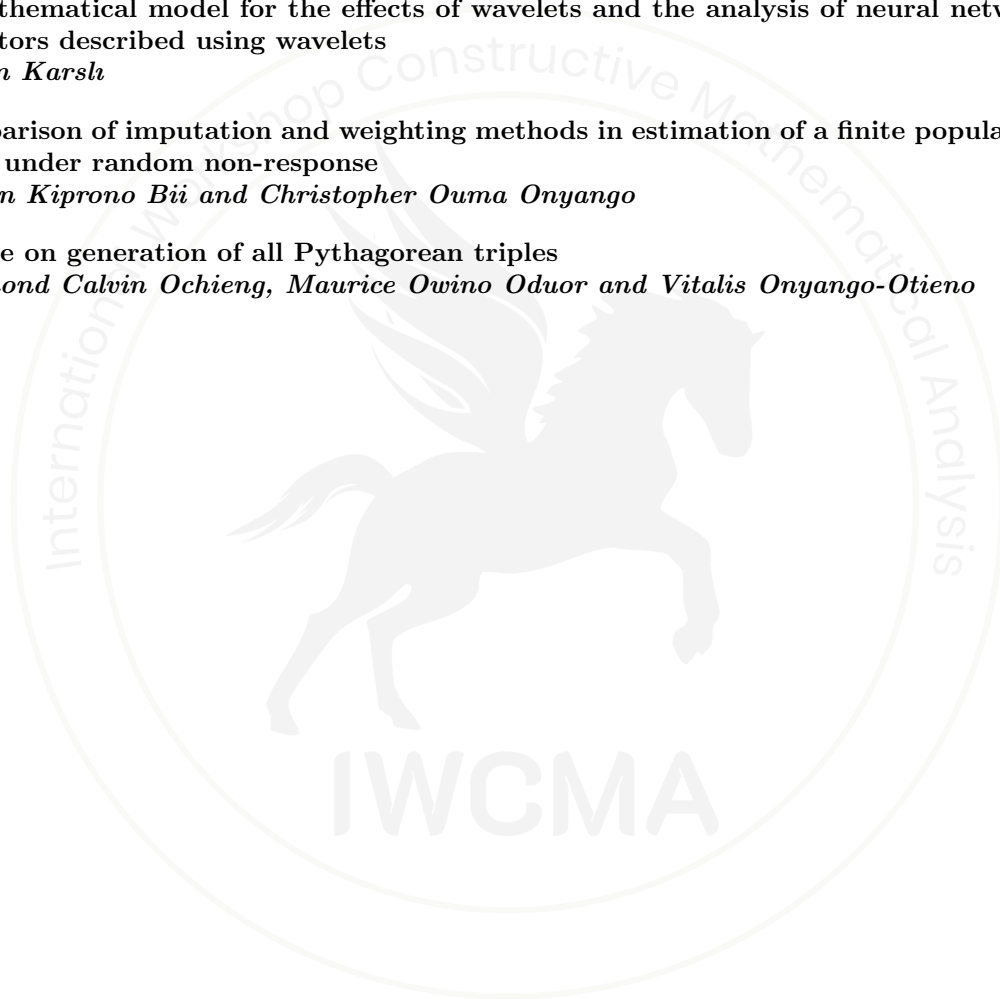
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## Some properties of squared Chlodovsky operators

Hüseyin Erhan Altın<sup>\*1</sup> and Harun Karsh<sup>2</sup>

ABSTRACT. The main purpose of this paper is to define a sequence of positive linear operators by means of the squared Chlodovsky basis functions. As a consequence of some certain inequalities we state that the second central moments of the constructed operators are smaller than the corresponding ones of the classical Bernstein-Chlodovsky operators. Furthermore, we estimate the rate of convergence in terms of the modulus of continuity and the class of Lipschitz functions.

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KEYWORDS: Chlodovsky operators, Linear positive operators, Rate of convergence

### 1. INTRODUCTION

The basis of the theory of approximation is the theorem discovered by Weierstrass in 1885 and the first constructive proof of this theorem was given by Bernstein [3] in 1912. He introduced a sequence of polynomials  $B_m : C[0, 1] \rightarrow C[0, 1]$  defined by

$$(B_m f)(x) = \sum_{k=0}^m f\left(\frac{k}{m}\right) p_{m,k}(x), \quad x \in [0, 1],$$

where Bernstein basis function is given by

$$(1) \quad p_{m,k}(x) = \binom{m}{k} x^k (1-x)^{m-k}, \quad m \in \mathbb{N}.$$

Later it was discovered that Bernstein polynomials have numerous significant properties, so new applications and generalizations are being found of it. The aim of these generalizations is to provide appropriate and powerful tools to application areas. One of these generalizations is the classical Bernstein-Chlodovsky operators are defined for a function  $f$  defined on  $[0, \infty)$  and bounded on every finite interval  $[0, b] \subset [0, \infty)$  by

$$(2) \quad (C_m f)(x) = \sum_{k=0}^m f\left(\frac{b_m k}{m}\right) p_{m,k}\left(\frac{x}{b_m}\right)$$

where  $p_{m,k}$  denotes as usual as (1) and  $(b_m)_{m=1}^{\infty}$  is a positive increasing sequence of reals with the properties

$$(3) \quad \lim_{m \rightarrow \infty} b_m = \infty, \quad \lim_{m \rightarrow \infty} \frac{b_m}{m} = 0.$$

These operators were introduced by Chlodovsky [4] in 1937 as a generalization of the Bernstein polynomials  $(B_m f)(x)$  on an infinite interval. Although there are many works on these operators, readers can find some results, collectively, on Chlodovsky operators in [9].

We can say that the last century has been very productive by means of working with linear positive operators. However working on squared type operators has become popular at the last decade. Now we give a look some remarkable studies on this topic. In [5], the authors obtained a new representation of the sum of the squared Bernstein polynomials and used it to validate a conjecture asserting that this sum is a convex function. They extended the result to some other classical approximation operators; Mirakjan-Favard-Szász operators, Meyer-König and Zeller operators and King-type operators. In [6], the authors provided some estimates of the second central moment of the squared Bernstein polynomials and also estimated the rate of approximation in terms of the modulus of continuity. In [7] and [8], the author obtained the Voronovskaya formula for sequence of positive linear operators constructed using the squared Bernstein polynomials and studied some approximation properties of positive linear operators defined by means of the powered Baskakov basis, respectively. In

[1] and [2], the author(s) derived a complete asymptotic expansion for a sequence of positive linear approximation operators defined by means of the squared Bernstein basis polynomials, Favard-Szász-Mirakjan fundamental functions and Baskakov fundamental functions. Also they studied the asymptotic properties of operators defined by means of squared Meyer-König and Zeller fundamental functions. In [10], the authors introduced a family of neural networks of multivariate square rational Bernstein operators defined by extending the artificial neural networks multivariate Bernstein by using square Bernstein polynomials and they studied the behavior of this neural network. And in [11], the authors defined a sequence of positive linear operators by means of the squared Szász-Mirakjan basis functions and estimated the rate of convergence in terms of the modulus of continuity and the class of Lipschitz functions. Furthermore, they showed the comparison and convergence of these operators with the help of some illustrative graphics.

In the light of these works, we consider the rational functions,

$$(4) \quad c_{m,k} = \frac{p_{m,k}^2}{\sum_{i=0}^m p_{m,i}^2}, \quad k = 0, 1, \dots, m,$$

and define the positive linear Bernstein-Chlodovsky type rational operators  $(SC_m f)$  by

$$(5) \quad (SC_m f)(x) = \sum_{k=0}^m f\left(\frac{b_m - k}{m}\right) c_{m,k} \left(\frac{x}{b_m}\right)$$

where  $c_{m,k}$  defined by (4) and  $(b_m)_{m=1}^{\infty}$  is a positive increasing sequence of reals with the properties (3). This work is organized as follows. In the next section, we give some auxiliary results. And in the final section, we prove the main results which are related with the second central moments and rate of convergence of these operators.

## 2. AUXILIARY RESULTS

In order to prove the main results, the following function will be the essential tool. For  $m \in \mathbb{N}$  and  $u \in [0, 1]$  define

$$(6) \quad g_m(u) = \frac{\int_0^1 \frac{ut(1-ut)^{m-1}}{\sqrt{t(1-t)}} dt}{\int_0^1 \frac{(1-ut)^m}{\sqrt{t(1-t)}} dt}.$$

Integration by parts shows that the integrals in (6) are differentiable with respect to the parameter  $u$ . Also define

$$(7) \quad M_m = \sup_{u \in [0,1]} g_m(u), \quad m = 1, 2, \dots$$

**Lemma 2.1.** *The sum of squared Bernstein-Chlodovsky basis function satisfies the following equality*

$$\sum_{k=0}^m p_{m,k}^2 \left(\frac{x}{b_m}\right) = \frac{1}{\pi} \int_0^1 \frac{\left(1 - 4\frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) t\right)^m}{\sqrt{t(1-t)}} dt$$

for  $x \in [0, b_m]$  and  $m \in \mathbb{N}$ .

*Proof.* Since

$$\sum_{k=0}^m p_{m,k} \left(\frac{x}{b_m}\right) e^{ik\theta} = \left(\frac{x}{b_m} e^{i\theta} + 1 - \frac{x}{b_m}\right)^m,$$

by Parseval's identity, we have

$$\begin{aligned} \sum_{k=0}^m p_{m,k}^2 \left(\frac{x}{b_m}\right) &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \left| \frac{x}{b_m} e^{i\theta} + 1 - \frac{x}{b_m} \right|^{2m} d\theta \\ &= \frac{2}{\pi} \int_0^{\pi/2} \left(1 - 4\frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) \sin^2 \theta\right)^m d\theta, \end{aligned}$$



and with  $\sin^2 \theta = t$ , we have

$$(8) \quad \sum_{k=0}^m p_{m,k}^2 \left( \frac{x}{b_m} \right) = \frac{1}{\pi} \int_0^1 \frac{\left(1 - 4 \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) t\right)^m}{\sqrt{t(1-t)}} dt.$$

□

**Lemma 2.2.** *The second central moment of the operators  $SC_m$  defined by (5) is given by*

$$(SC_m) \left( (t-x)^2; x \right) = \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) g_m(u)$$

with  $u = 4 \left( \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) \right) \in [0, 1]$  and  $g_m(u)$  is defined by equation (6).

*Proof.* Starting with the equality

$$(9) \quad \sum_{k=0}^m p_{m,k} \left( \frac{x}{b_m} \right) e^{u \left( \frac{x}{b_m} - \frac{k}{m} \right)} e^{ik\theta} = e^{\frac{ux}{b_m}} \left( \frac{x}{b_m} e^{-\frac{u}{m}} e^{i\theta} + 1 - \frac{x}{b_m} \right)^m$$

and differentiating both sides of (9) with respect to  $u$ , and taking  $u = 0$ , we obtain

$$\sum_{k=0}^m p_{m,k} \left( \frac{x}{b_m} \right) \left( \frac{x}{b_m} - \frac{k}{m} \right) e^{ik\theta} = \frac{x}{b_m} \left( \frac{x}{b_m} e^{i\theta} + 1 - \frac{x}{b_m} \right)^{m-1} \left( \frac{x}{b_m} e^{i\theta} - e^{i\theta} + 1 - \frac{x}{b_m} \right).$$

Using Parseval's equality, we have

$$(10) \quad \sum_{k=0}^m p_{m,k}^2 \left( \frac{x}{b_m} \right) \left( \frac{x}{b_m} - \frac{k}{m} \right)^2 = \frac{8 \left( \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) \right)^2}{\pi} \int_0^{\pi/2} \left(1 - 4 \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) \sin^2 \theta\right)^{m-1} \sin^2 \theta d\theta.$$

Here, taking  $\sin^2 \theta = t$  in (10), we obtain

$$(11) \quad \sum_{k=0}^m p_{m,k}^2 \left( \frac{x}{b_m} \right) \left( \frac{x}{b_m} - \frac{k}{m} \right)^2 = \frac{4 \left( \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) \right)^2}{\pi} \int_0^1 \frac{\left(1 - 4 \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) t\right)^{m-1}}{\sqrt{t(1-t)}} t dt.$$

With  $u = 4 \left( \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) \right) \in [0, 1]$ , and from (5), (11) with (8), we have

$$(12) \quad (SC_m) \left( (t-x)^2; x \right) = \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) g_m(u)$$

which completes the proof. □

### 3. MAIN RESULTS

**Theorem 3.1.** *The following inequality*

$$(13) \quad \frac{1}{2b_m^2} \frac{x(b_m - x)}{m} \leq (SC_m) \left( (t-x)^2; x \right),$$

is valid for all  $x \in [0, b_m]$  and  $m \geq 1$ .

*Proof.* Since  $\frac{1}{\sqrt{t(1-t)}} = (2 \arcsin \sqrt{t})'$ , using integration by parts, we obtain

$$(14) \quad \begin{aligned} \int_0^1 \frac{(1-ut)^m}{\sqrt{t(1-t)}} dt &= 2 \int_0^1 (1-ut)^m \left( \arcsin \sqrt{t} \right)' dt \\ &= (1-u)^m \pi + 2mu \int_0^1 (1-ut)^{m-1} \arcsin \sqrt{t} dt \\ &= (\varepsilon_m(u) + 1) 2mu \int_0^1 (1-ut)^{m-1} \arcsin \sqrt{t} dt \end{aligned}$$

where

$$\begin{aligned}
 \varepsilon_m(u) &= \frac{(1-u)^m \pi}{2mu \int_0^1 (1-ut)^{m-1} \arcsin \sqrt{t} dt} \\
 &< \frac{(1-u)^m \pi}{2mu \int_0^1 (1-u)^{m-1} \arcsin \sqrt{t} dt} \\
 (15) \quad &= \frac{2(1-u)}{mu},
 \end{aligned}$$

for  $u \in [0, 1]$  and  $m \geq 1$ . From (14), using the inequality

$$\arcsin \sqrt{t} \leq \frac{t}{\sqrt{t(1-t)}}, \quad t \in [0, 1],$$

we obtain

$$(16) \quad \int_0^1 \frac{(1-ut)^m}{\sqrt{t(1-t)}} dt \leq (\varepsilon_m(u) + 1) 2mu \int_0^1 \frac{t(1-ut)^{m-1}}{\sqrt{t(1-t)}} dt.$$

From (16) and (15), we deduce that:

$$mg_m(u) \geq \frac{1}{2(\varepsilon_m(u) + 1)} \geq \frac{1}{2\left(\frac{2(1-u)}{mu} + 1\right)}$$

hence

$$mg_m(u) \geq \frac{1}{2} - \frac{1-u}{2+(n-2)u} = \frac{1}{2} - \frac{1-4\left(1-\frac{x}{b_m}\right)\frac{x}{b_m}}{2\left(1+2(n-2)\left(1-\frac{x}{b_m}\right)\frac{x}{b_m}\right)},$$

with  $u := 4\left(\frac{x}{b_m}\left(1-\frac{x}{b_m}\right)\right) \in [0, 1]$  and  $m = 1, 2, \dots$ . Finally, in virtue of (12), we obtain:

$$\begin{aligned}
 m(SC_m)\left((t-x)^2; x\right) &= \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) mg_m(u) \\
 &\geq \frac{\frac{x}{b_m} \left(1 - \frac{x}{b_m}\right)}{2} - \frac{\frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) \left(1 - 4\left(1 - \frac{x}{b_m}\right) \frac{x}{b_m}\right)}{2\left(1 + 2(n-2)\left(1 - \frac{x}{b_m}\right) \frac{x}{b_m}\right)}
 \end{aligned}$$

and the proof is complete.  $\square$

**Corollary 3.2.** *The second central moments of the operators  $SC_m$  defined by (5) satisfy the following inequalities:*

$$\begin{aligned}
 (17) \quad \frac{1}{2b_m^2} \frac{x(b_m-x)}{m} &\leq (SC_m)\left((t-x)^2; x\right) \leq p M_p \frac{x(b_m-x)}{m}, \\
 (SC_m)\left((t-x)^2; x\right) &\leq p M_p (C_m)\left((t-x)^2; x\right)
 \end{aligned}$$

for  $x \in [0, b_m]$  and  $1 \leq p \leq m$ .

*Proof.* Using (12) and (7), we have

$$\begin{aligned}
 (18) \quad (SC_m)\left((t-x)^2; x\right) &= \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) g_m(u) \\
 &\leq M_m \frac{x}{b_m} \left(1 - \frac{x}{b_m}\right) \\
 &\leq p M_p \frac{x(b_m-x)}{m} \\
 &\leq p M_p (C_m)\left((t-x)^2; x\right)
 \end{aligned}$$

for all  $1 \leq p \leq m$ . Putting together (13) and (18), we get the desired result which states that the second central moments  $(SC_m)\left((t-x)^2; x\right)$  of the operators  $SC_m$  are of order exactly

$\frac{x(b_m-x)}{m}$ , and are smaller than the corresponding ones of the classical Bernstein-Chlodovsky operators.  $\square$

In the approximation of a function by positive linear operators not only the convergence of operators is required but also the speed of convergence is important. The rate of convergence depends on the smoothness properties of the function and appropriate tool for estimating the smoothness of function are represented by modulus of continuity. We compute the rate of convergence of the constructed operators in terms of modulus of continuity and class of Lipschitz function.

**Theorem 3.3.** *For any  $f \in C_B[0, \infty)$  (space of all bounded and continuous functions on  $[0, \infty)$ ), we have the estimate*

$$|(SC_m)(f; x) - f(x)| \leq \left(1 + \sqrt{p M_p}\right) \omega\left(f; \sqrt{\frac{x(b_m-x)}{m}}\right)$$

where  $1 \leq p \leq m$ ,  $M_p$  is given by (7) and  $\omega(f; \cdot)$  is the modulus of continuity.

*Proof.* Since

$$\begin{aligned} |(SC_m)(f; x) - f(x)| &= \left| \sum_{k=0}^m f\left(\frac{b_m k}{m}\right) c_{m,k}\left(\frac{x}{b_m}\right) - f(x) \right| \\ &\leq \sum_{k=0}^m c_{m,k}\left(\frac{x}{b_m}\right) \left| f\left(\frac{b_m k}{m}\right) - f(x) \right|, \end{aligned}$$

by using Cauchy-Schwartz inequality and the property

$$|f(t) - f(x)| \leq \omega(f; \delta) \left(1 + \frac{|t-x|}{\delta}\right),$$

we get

$$\begin{aligned} |(SC_m)(f; x) - f(x)| &\leq \sum_{k=0}^m c_{m,k}\left(\frac{x}{b_m}\right) \left\{1 + \frac{1}{\delta} \left| \frac{b_m k}{m} - x \right| \right\} \omega(f; \delta) \\ &\leq \left\{1 + \frac{1}{\delta} \left( \sum_{k=0}^m c_{m,k}\left(\frac{x}{b_m}\right) \left(\frac{b_m k}{m} - x\right)^2 \right)^{1/2} \left( \sum_{k=0}^m c_{m,k}\left(\frac{x}{b_m}\right) \right)^{1/2} \right\} \omega(f; \delta) \\ &\leq \left\{1 + \frac{1}{\delta} \left( (SC_m)\left((\cdot - x)^2; x\right) \right)^{1/2} \left( (SC_m)(1; x) \right)^{1/2} \right\} \omega(f; \delta). \end{aligned}$$

Using the fact that  $(SC_m)(1; x) = 1$  and (17), we have

$$|(SC_m)(f; x) - f(x)| \leq \left\{1 + \frac{1}{\delta} \left( p M_p \frac{x(b_m-x)}{m} \right)^{1/2} \right\} \omega(f; \delta),$$

choosing  $\delta = \sqrt{\frac{x(b_m-x)}{m}}$ , we obtain the result.  $\square$

**Corollary 3.4.** *For any  $f \in C_B[0, \infty)$ , the sequence of operators  $(SC_m f)_{m \in \mathbb{N}}$  converges uniformly to  $f$  on  $[0, \infty)$ .*

**Theorem 3.5.** *Let  $f \in C_B[0, \infty)$ ,  $M > 0$  and  $0 < \mu \leq 1$ , then*

$$|(SC_m)(f; x) - f(x)| \leq M (\delta_m(x))^{\frac{\mu}{2}}$$

for each  $f \in Lip_M(\mu) = \{f : |f(\eta_1) - f(\eta_2)| \leq M |\eta_1 - \eta_2|^\mu, \eta_1, \eta_2 \in [0, \infty)\}$  and for  $\delta_m(x) = (SC_m)\left((t-x)^2; x\right)$

*Proof.* We prove this theorem by using the definition of Lipschitz function and Hölder's inequality.

$$\begin{aligned} |(SC_m)(f; x) - f(x)| &\leq SC_m(|f(t) - f(x)|; x) \\ &\leq M SC_m(|t-x|^\mu; x). \end{aligned}$$

Therefore

$$\begin{aligned}
 |(SC_m)(f; x) - f(x)| &\leq M \sum_{k=0}^m c_{m,k} \left( \frac{x}{b_m} \right) \left| \frac{b_m}{m} k - x \right|^\mu \\
 &\leq M \sum_{k=0}^m \left( c_{m,k} \left( \frac{x}{b_m} \right) \right)^{\frac{2-\mu}{2}} \left( c_{m,k} \left( \frac{x}{b_m} \right) \right)^{\frac{\mu}{2}} \left| \frac{b_m}{m} k - x \right|^\mu \\
 &\leq M \left\{ \left( \sum_{k=0}^m c_{m,k} \left( \frac{x}{b_m} \right) \right)^{\frac{2-\mu}{2}} \left( \sum_{k=0}^m c_{m,k} \left( \frac{x}{b_m} \right) \left| \frac{b_m}{m} k - x \right|^2 \right)^{\frac{\mu}{2}} \right\} \\
 &= M \left( (SC_m) \left( (\cdot - x)^2; x \right) \right)^{\frac{\mu}{2}}.
 \end{aligned}$$

Choosing  $\delta_m(x) = (SC_m) \left( (\cdot - x)^2; x \right)$ , proof is completed.  $\square$

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## On new means generated by inverse of eigenfunctions of $(p, q)$ -Laplacian

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ABSTRACT. In this paper, we generalize Neuman-Sándor mean  $M$  and Seiffert mean  $P$  as an application of the inverse of eigenfunctions generalized hyperbolic functions with two parameters. Moreover, two-sided inequalities involving these generalized means are established.

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### 1. INTRODUCTION

In 1879, Lindberg [17] originally introduced generalized trigonometric functions. In 1995, Lindqvist [18] also introduced these functions, his work was highly cited by several authors. The recent literature on these functions includes several dozens of papers. These functions were used to study problems of existence, bifurcation and oscillation of solutions of differential equations, applications to differential equations involving the  $p$ -Laplacian, simple generalization of the classical trigonometric and hyperbolic functions, generalization of elliptic integrals of the first and the second kind, and generalization of means of two variables. The reader is referred to see [5, 7, 6, 11, 14, 27, 29, 31] and the bibliography therein.

The *Gaussian hypergeometric function* is defined by

$$F(a, b; c; z) = {}_2F_1(a, b; c; z) = \sum_{n=0}^{\infty} \frac{(a, n)(b, n)}{(c, n)} \frac{z^n}{n!}, \quad |z| < 1,$$

where  $(a, n)$  denotes the shifted factorial function

$$(a, n) = a(a+1)(a+2)\dots(a+n-1), \quad n = 1, 2, 3, \dots,$$

and  $(a, 0) = 1$  for  $a \neq 1$ . The reader is referred to see [4] for the applications of Gaussian hypergeometric function in various fields of the mathematical and natural sciences.

Special functions, such as classical *gamma function*  $\Gamma$ , the *digamma function*  $\psi$  and the *beta function*  $B(\cdot, \cdot)$  have close relation with hypergeometric function. For  $x, y > 0$ , these functions are defined by

$$\Gamma(x) = \int_0^{\infty} e^{-t} t^{x-1} dt, \quad \psi(x) = \frac{\Gamma'(x)}{\Gamma(x)}, \quad B(x, y) = \frac{\Gamma(x)\Gamma(y)}{\Gamma(x+y)},$$

respectively. The hypergeometric function can be represented in the integral form as follows

$$(1) \quad F(a, b; c; z) = \frac{\Gamma(c)}{\Gamma(b)(c-b)} \int_0^1 t^{b-1} (1-t)^{c-b-1} (1-zt)^{-a} dt.$$

The eigenfunction  $\sin_p$  of the so-called one-dimensional  $p$ -Laplacian problem [15]

$$-\Delta_p u = -(|u'|^{p-2} u')' = \lambda |u|^{p-2} u, \quad u(0) = u(1) = 0, \quad p > 1,$$

is the inverse function of  $F_p : [0, 1] \rightarrow [0, \frac{\pi_p}{2}]$ , defined as

$$F_p(x) = \arcsin_p(x) = \int_0^x (1-t^p)^{-\frac{1}{p}} dt,$$

where

$$\pi_p = 2\arcsin_p(1) = \frac{2}{p} \int_0^1 (1-s)^{-\frac{1}{p}} s^{\frac{1}{p}-1} ds = \frac{2}{p} B\left(1 - \frac{1}{p}, \frac{1}{p}\right) = \frac{2\pi}{p \sin\left(\frac{\pi}{p}\right)}.$$



The function  $\arcsin_p$  is called the generalized inverse sine function, and its inverse function  $\sin_p : [0, \pi_p/2] \rightarrow [0, 1]$  is called generalized sine function. For  $x \in [\pi_p/2, \pi_p]$ , one can extend the function  $\sin_p$  to  $[0, \pi_p]$  by defining  $\sin_p(x) = \sin(\pi_p - x)$ , and further extension can be achieved on  $\mathbb{R}$  by oddness and  $2\pi$ -periodicity. The range of  $p$  is restricted to  $(1, \infty)$  because only in this case  $\sin_p(x)$  can be made periodic like usual sine function.

Similarly, the other generalized inverse trigonometric and hyperbolic functions  $\arccos_p : (-1, 1) \rightarrow (-a_p, a_p)$ ,  $\arctan_p : (-\infty, \infty) \rightarrow (-a_p, a_p)$ ,  $\operatorname{arcsinh}_p : (-\infty, \infty) \rightarrow (-\infty, \infty)$ ,  $\operatorname{arctanh}_p : (-1, 1) \rightarrow (-\infty, \infty)$  are defined as follows

$$(2) \quad \begin{aligned} \arccos_p(x) &= \int_0^{(1-x^p)^{\frac{1}{p}}} (1 - |t|^p)^{-\frac{1}{p}} dt, & \arctan_p(x) &= \int_0^x (1 + |t|^p)^{-1} dt, \\ \operatorname{arcsinh}_p(x) &= \int_0^x (1 + |t|^p)^{-\frac{1}{p}} dt, & \operatorname{arctanh}_p(x) &= \int_0^x (1 - |t|^p)^{-1} dt, \end{aligned}$$

where  $a_p = \pi_p/2$ . Above inverse generalized trigonometric and hyperbolic functions coincide with usual trigonometric and hyperbolic functions for  $p = 2$ .

Generalization of means with two parameters: For two positive real numbers  $a$  and  $b$ , we define arithmetic mean  $A$ , geometric mean  $G$ , logarithmic mean  $L$ , two Seiffert means  $P$  and  $T$ , and Neuman-Sándor mean  $M$  introduced in [21] as follows,

$$A = A(a, b) = \frac{a + b}{2}, \quad G = G(a, b) = \sqrt{ab},$$

$$L = L(a, b) = \frac{a - b}{\log(a) - \log(b)}, \quad a \neq b,$$

$$P = P(a, b) = \frac{a - b}{2\arcsin\left(\frac{a-b}{a+b}\right)},$$

$$T = T(a, b) = \frac{a - b}{2\arctan\left(\frac{a-b}{a+b}\right)},$$

$$M = M(a, b) = \frac{a - b}{2\operatorname{arcsinh}\left(\frac{a-b}{a+b}\right)}.$$

Neuman [23, 24] generalized the logarithmic mean  $L$ , two Seiffert means  $P$  and  $T$ , and the Neuman-Sándor mean  $M$  by introducing the the  $p$ -version of the Schwab-Borchardt mean  $SB_p$  as follows

$$L_p = L_p(a, b) = SB_p(A_{p/2}, G) = \frac{A_{p/2}v_p}{\operatorname{arctanh}_p(v_p)},$$

$$P_p = P_p(a, b) = SB_p(G, A_{p/2}) = \frac{A_{p/2}v_p}{\arcsin_p(v_p)},$$

$$T_p = T_p(a, b) = SB_p(A_{p/2}, A_p) = \frac{A_{p/2}v_p}{\arctan_p(v_p)},$$

$$M_p = M_p(a, b) = SB_p(A_p, A_{p/2}) = \frac{A_{p/2}v_p}{\operatorname{arcsinh}_p(v_p)},$$

where

$$SB_p(a, b) = b F\left(\frac{1}{p}, \frac{1}{p}; 1 + \frac{1}{p}, 1 - \left(\frac{a}{b}\right)^p\right)^{-1},$$

$$v_p = \frac{|x^{p/2} - y^{p/2}|}{x^{p/2} + y^{p/2}},$$

and  $A_p = A_p(a, b)$  is a power mean of order  $p$ .

Motivated by the work of Neuman, Bhayo and Sándor gave a natural and new generalization of  $L$ ,  $P$ ,  $T$  and  $M$  in [10, Theorem 2.1] by utilizing the generalized trigonometric and generalized hyperbolic functions as follows:

For  $p \geq 2$  and  $a > b > 0$ , the following functions define a mean of two variables  $a$  and  $b$

$$\begin{aligned}
 \tilde{P}_p = \tilde{P}_p(a, b) &= \frac{a-b}{2\arcsin_p\left(\frac{a-b}{a+b}\right)} = \frac{x}{\arcsin_p(x)} A = \frac{A}{F\left(\frac{1}{p}, \frac{1}{p}; 1 + \frac{1}{p}; x^p\right)}, \\
 \tilde{T}_p = \tilde{T}_p(a, b) &= \frac{a-b}{2\arctan_p\left(\frac{a-b}{a+b}\right)} = \frac{x}{\arctan_p(x)} A = \frac{A \cdot (1+x^p)^{1/p}}{F\left(\frac{1}{p}, \frac{1}{p}; 1 + \frac{1}{p}; \frac{x^p}{1+x^p}\right)}, \\
 \tilde{L}_p = \tilde{L}_p(a, b) &= \frac{a-b}{2\operatorname{artanh}_p\left(\frac{a-b}{a+b}\right)} = \frac{x}{\operatorname{artanh}_p(x)} A = \frac{A}{F\left(1, \frac{1}{p}; 1 + \frac{1}{p}; x^p\right)}, \\
 \tilde{M}_p = \tilde{M}_p(a, b) &= \frac{a-b}{2\operatorname{arsinh}_p\left(\frac{a-b}{a+b}\right)} = \frac{x}{\operatorname{arsinh}_p(x)} A = \frac{A \cdot (1+x^p)^{1/p}}{F\left(1, \frac{1}{p}; 1 + \frac{1}{p}; \frac{x^p}{1+x^p}\right)},
 \end{aligned}
 \tag{3}$$

where  $x = (a-b)/(a+b)$ .

In [27], generalized inverse trigonometric function  $\arcsin_p$  was extended for two parameters  $p, q > 1$  as follows,

$$\arcsin_{p,q}(x) = \int_0^x (1-t^q)^{-\frac{1}{p}} dt,$$

for  $x \in (0, 1)$ .

Letting  $t = z^{1/q}$ , we observe that

$$\arcsin_{p,q}(x) = \frac{1}{q} \int_0^{x^q} z^{1/q-1} (1-z)^{-1/p} dz = \frac{1}{q} \tilde{B}\left(\frac{1}{q}, 1 - \frac{1}{p}, x^q\right),$$

where  $\tilde{B}(a, b, x)$  is incomplete beta function defined as

$$\tilde{B}(a, b, x) = \int_0^x t^{a-1} (1-t)^{b-1} dt.$$

The function  $\arcsin_{p,q}(x)$  is the inverse function of  $\sin_{p,q}$ , defined on the the interval  $[0, \pi_{p,q}/2]$ , where

$$\pi_{p,q} = 2\arcsin_{p,q}(1) = \frac{2}{q} B\left(1 - \frac{1}{p}, \frac{1}{q}\right) = \frac{2\pi}{q \sin\left(\frac{\pi}{p}\right)}.$$

For  $T = \pi_{p,q}$  the function  $u(t) = \sin_{p,q}(t)$  is a solution to the following problem considered by Drábek and Manásevich Let  $\phi_p(x) = |x|^{p-2}x$ . For  $T, \lambda > 0$  and  $p, q > 1$

$$\begin{cases} (\phi_p(u'))' + \lambda \phi_q(u) = 0, & t \in (0, T), \\ u(0) = u(T) = 0. \end{cases}$$

For  $x \in (0, 1)$ , we also define  $\arccos_{p,q}(x) = \arcsin_{p,q}((1-x^p)^{1/q})$ , and

$$\cos_{p,q}(x) = \frac{d}{dx} \sin_{p,q}(x), \quad x \in [0, \pi_{p,q}/2].$$

Letting  $y = \sin_{p,q}(x)$ , we get  $\cos_{p,q}(x) = (1 - (\sin(x)^q))^{1/p}$  and

$$|\cos_{p,q}|^p + |\sin_{p,q}|^q = 1.$$

Similarly, the generalized tangent function is defined as  $\tan_{p,q} = (\sin_{p,q}(x))/(\cos_{p,q}(x))$ , and its inverse is denoted by  $\arctan_{p,q}$ , see [9].

Motivated by the work of Takeuchi [27], Bhayo and Vuorinen [11] defined the generalized inverse hyperbolic sine function  $\operatorname{arsinh}_{p,q}$  as follows

$$\operatorname{arsinh}_{p,q}(x) = \int_0^x (1+t^q)^{-\frac{1}{p}} dt = x F\left(\frac{1}{p}, \frac{1}{q}; 1 + \frac{1}{q}; x^q\right),$$

for  $x \in (0, 1)$ . In similar fashion, generalized hyperbolic cosine and tangent function are defined by

$$\cosh_{p,q}(x) = \frac{d}{dx} \sinh_{p,q}(x), \quad \tanh_{p,q}(x) = \frac{\sinh_{p,q}(x)}{\cosh_{p,q}(x)}, \quad x \geq 0,$$

and the inverse of  $\tanh_{p,q}$  is denoted by  $\operatorname{artanh}$ . It follows from the definition that

$$|\cosh_{p,q}|^p - |\sinh_{p,q}|^q = 1.$$

In [11], it was proved that for  $p, q > 1$  and  $x \in (0, 1)$  one has

$$(4) \quad x \left( 1 + \frac{x^q}{p(1+q)} \right) < \arcsin_{p,q}(x) < \min \left\{ \frac{\pi_{p,q}}{2}, (1-x^q)^{-1/(p(1+q))} x \right\}$$

$$(5) \quad \left( \frac{x^p}{1+x^q} \right)^{1/p} r_{p,q}(x) < \operatorname{arcsinh}_{p,q}(x) < \left( \frac{x^p}{1+x^q} \right)^{1/p} s_{p,q}(x),$$

where

$$r_{p,q}(x) = \max \left\{ \left( 1 - \frac{qx^q}{p(1+q)(1+x^q)} \right)^{-1}, (1+x^q)^{1/p} \left( \frac{pq+p+qx^q}{p(1+q)} \right)^{-1/q} \right\}$$

and  $s_{p,q}(x) = (1-x^p/(1+x^q))^{-q/(p(1+q))}$ .

It is easy to observe that  $(p, q)$ -functions reduce to  $p$ -functions for  $p = q$ , and  $\pi_{p,q}$  to  $\pi_p$ . It follows from definition that  $\pi_{p,q} \leq \pi_p$  for  $q \geq p > 1$ .

### 1.1. Preliminaries.

**Lemma 1.1.** [3, Theorem 2] For  $-\infty < a < b < \infty$ , let  $f, g : [a, b] \rightarrow \mathbb{R}$  be continuous on  $[a, b]$ , and be differentiable on  $(a, b)$ . Let  $g'(x) \neq 0$  on  $(a, b)$ . If  $f'(x)/g'(x)$  is increasing (decreasing) on  $(a, b)$ , then so are

$$\frac{f(x) - f(a)}{g(x) - g(a)} \quad \text{and} \quad \frac{f(x) - f(b)}{g(x) - g(b)}.$$

If  $f'(x)/g'(x)$  is strictly monotone, then the monotonicity in the conclusion is also strict.

**Lemma 1.2.** [9] For  $p, q > 1$ , the following inequalities hold true,

- (1)  $\frac{x}{\arcsin_{p,q}(x)} < \frac{\sin_{p,q}(x)}{x}, \quad x \in (0, 1),$
- (2)  $\frac{x}{\operatorname{arcsinh}_{p,q}(x)} < \frac{\sinh_{p,q}(x)}{x}, \quad x \in (0, \infty),$
- (3)  $\frac{x}{\arctan_{p,q}(x)} < \frac{\tan_{p,q}(x)}{x}, \quad x \in (0, 1),$
- (4)  $\frac{x}{\operatorname{arctanh}_{p,q}(x)} < \frac{\tanh_{p,q}(x)}{x}, \quad x \in (0, 1).$

For easy reference we recall some well-known inequalities from the literature as follows.

**Cauchy-Bouniakowski inequality.** If  $f, g : [a, b] \rightarrow \mathbb{R}$  are integrable, then

$$(6) \quad \left( \int_a^b f(x)g(x)dx \right)^2 \leq \int_a^b f(x)^2 dx \int_a^b g(x)^2 dx.$$

**Pólya-Szegő inequality.** If  $f, g : [a, b] \rightarrow \mathbb{R}$  are integrable, and for all  $x \in [a, b]$

$$0 < \alpha < f(x) < A, \quad 0 < \beta < g(x) < B,$$

then

$$(7) \quad \frac{\int_a^b f(x)^2 dx \int_a^b g(x)^2 dx}{\left( \int_a^b f(x)g(x)dx \right)^2} \leq K(\alpha, A, \beta, B),$$

where

$$K = K(\alpha, A, \beta, B) = \frac{1}{4} \left( \sqrt{\frac{AB}{\alpha\beta}} + \sqrt{\frac{\alpha\beta}{AB}} \right)^2.$$

**Minkowski's inequality.** Let  $f, g : [a, b] \rightarrow \mathbb{R}$  be integrable and  $f, g > 0$ . Write

$$h_t(f) = \left( \int_a^b f(x)^t dx \right)^{1/t}.$$

Then one has

$$(8) \quad h_t(f+g) \leq h_t(f) + h_t(g), \quad \text{for } t \geq 1,$$

$$(9) \quad h_t(f+g) \geq h_t(f) + h_t(g), \quad \text{for } t \leq 1.$$

**Diaz-Metcalf inequality.** Let  $f, g : [a, b] \rightarrow \mathbb{R}$  be integrable and suppose that there exist constants  $m$  and  $M$  such that

$$m \leq g(x)/f(x) \leq M.$$

Then one has

$$(10) \quad \int_a^b g^2(x)dx + m \cdot M \cdot \int_a^b f^2(x)dx \leq (m + M) \cdot \int_a^b f(x)g(x)dx.$$

The paper is organized as follows. In section 1, we give the definition of the special functions involved in our formulation of the main results, and definition of generalized means with one parameter and the statement their previous results. Few lemmas and some well-known inequalities are given in the subsection ('preliminaries'). In section 2, we give the definition of the generalized means with two parameters and the statement of the main results in the form of theorems. Section 3 is consisting of the proof of theorems. In Section 4, we present a conjecture.

## 2. MAIN RESULTS

The mean  $\tilde{P}_p$  and  $\tilde{M}_p$  can be further generalized for two parameters  $p, q \geq 2$  as follows,

$$(11) \quad \begin{aligned} \tilde{P}_{p,q} = \tilde{P}_{p,q}(a, b) &= \frac{a - b}{2\arcsin_p\left(\frac{a-b}{a+b}\right)} = \frac{x}{\arcsin_{p,q}(x)}A, \\ \tilde{M}_{p,q} = \tilde{M}_{p,q}(a, b) &= \frac{a - b}{2\operatorname{arsinh}_p\left(\frac{a-b}{a+b}\right)} = \frac{x}{\operatorname{arsinh}_{p,q}(x)}A \\ \tilde{T}_{p,q} = \tilde{T}_p(a, b) &= \frac{a - b}{2\arctan_{p,q}\left(\frac{a-b}{a+b}\right)} = \frac{x}{\arctan_{p,q}(x)}A \\ \tilde{L}_{p,q} = \tilde{L}_p(a, b) &= \frac{a - b}{2\operatorname{artanh}_{p,q}\left(\frac{a-b}{a+b}\right)} = \frac{x}{\operatorname{artanh}_{p,q}(x)}A, \end{aligned}$$

where  $x = (a - b)/(a + b)$  with  $0 < b < a$ . Clearly, for  $p = q$  we have  $\tilde{P}_{p,p} = \tilde{P}_p$ ,  $\tilde{M}_{p,p} = M_p$ ,  $\tilde{T}_{p,p} = T_p$ , and  $\tilde{L}_{p,p} = L_p$ . For  $p = 2$ ,  $\tilde{P}_2 = P$ ,  $\tilde{T}_2 = T$ ,  $\tilde{L}_2 = L$ ,  $\tilde{M}_2 = M$ ,  $\tilde{T}_2 = T$ , and  $\tilde{L}_2 = M$ .

**Theorem 2.1.** For  $p, q \geq 2$  and  $a > b > 0$ , the function  $\tilde{P}_{p,q}$  and  $\tilde{M}_{p,q}$  define a mean function of two variables  $a$  and  $b$ .

**Theorem 2.2.** For  $2 \leq p < q$  and  $x = (a - b)/(a + b)$  with  $0 < b < a$ , we have

$$(12) \quad P_{p,q} M_{p,q} \leq (P_{2p,2q})^2 \leq k(x, p, q) P_{p,q} M_{p,q},$$

where

$$k(x, p, q) = \frac{((1 + x^q)^{2/p} + (1 - x^q)^{2/p})^2}{4((1 - x^{2q})^{1/p})}.$$

**Theorem 2.3.** For  $a > b > 0$  and  $x = (a - b)/(a + b)$ , we have

$$(13) \quad \frac{1}{\tilde{P}_{p,q}} + \frac{r}{\tilde{M}_{p,q}} \leq \frac{r(p, q, x) + 1}{\tilde{P}_{2p,2q}},$$

$$(14) \quad (\tilde{P}_{2p,2q})^{2p} \left( \frac{1}{(\tilde{P}_{p,q})^p} + \frac{1}{(\tilde{M}_{p,q})^p} \right) \leq R(p, q, x),$$

where

$$r(p, q, x) = \frac{(1 + x^q)^{1/(2p)}}{(1 - x^q)},$$

$$R(p, q, x) = [(1 - x^{2q})^{1/(2p)} + (1 - x^{2q})^{-1/(2p)}]^{1/(2p)} / 2^{2p-1}.$$

## 3. PROOF OF MAIN RESULTS

**Proof of Theorem 2.1.** Clearly, for  $t \in (0, 1)$  and  $2 < p < q$ , the following inequality

$$\int_0^x (1+t^q)^{-1/p} dt < \int_0^x (1-t^q)^{-1/p} dt$$

implies

$$(15) \quad \tilde{P}_{p,q} < \tilde{M}_{p,q}.$$

Similarly, for  $t \in (0, 1)$  and  $2 < p < q$  the inequality

$$\int_0^x (1+t^q)^{-1/p} dt < \int_0^x (1+t^p)^{-1/p} dt$$

and

$$\int_0^x (1-t^{-q})^{-1/p} dt > \int_0^x (1-t^{-p})^{-1/p} dt$$

yield

$$(16) \quad \tilde{M}_{p,q} > \tilde{M}_p \quad \text{and} \quad \tilde{P}_{p,q} < \tilde{P}_p.$$

Since  $\tilde{P}_{p,q}$  is a mean, by (15) to prove that  $\tilde{M}_{p,q}$  is a mean, it is sufficient to find an upper bound for  $\tilde{M}_{p,q}$ , which is a mean.

Let  $Q = Q(a, b) = \sqrt{(a^2 + b^2)/2}$ . Let  $(a - b)/(a + b) = x$ . Then it is easy to see that one has the following identity:

$$(17) \quad Q/A = \sqrt{1+x^2}.$$

Now, since

$$\frac{A}{\tilde{M}_{p,q}} = \frac{1}{x} \int_0^x (1+t^q)^{-1/p} dt.$$

Therefore, by (17), to prove that  $A/Q \leq A/\tilde{M}_{p,q}$ , it is sufficient to prove that

$$(18) \quad (1+t^q)^{1/q} \leq (1+x^2)^{1/2}.$$

For  $q \geq 2$  and  $t \in [0, 1]$ , we have

$$(1+t^q)^{1/p} \leq (1+t^2)^{1/p} \leq (1+x^2)^{1/p} \leq (1+x^2)^{1/2},$$

so (18) follows. This implies that  $A/\tilde{M}_{p,q} \geq A/Q$ , so

$$(19) \quad \tilde{M}_{p,q} \leq Q.$$

By (15) and (19) it follows that  $\tilde{M}_{p,q}$  is also a mean. □

**An other proof of  $\tilde{P}_{p,q}$  mean.** By definition, we have

$$\frac{\tilde{P}_{p,q}}{A} = \frac{2z}{\arcsin_{p,q}(z)},$$

where  $z = (x - y)/(x + y)$  and  $x > y$ . Now, from the inequalities

$$x < \arcsin_{p,q}(t) < \frac{\pi_{p,q}}{2}x,$$

we get the double inequality:

$$(20) \quad \frac{2}{\pi_{p,q}}A < \tilde{P}_{p,q} < A,$$

assume that  $\alpha_{p,q} \geq 2$ . Now, by Jordan's inequality  $\sin x > (2/\pi)x > x/2$ , we get  $\pi_{p,q} \leq \pi_p < 4$ . Let  $0 < y < x$ . Then from (20) we get

$$y < \tilde{P}_{p,q} < x,$$

where the right hand side is trivial, as  $A = (x + y)/2 < x$ , and for the left hand side,  $(2/\pi_{p,q})A > (4/\pi_{p,q})y > y$ , by  $\pi_{p,q} < 4$ . This implies that  $\tilde{P}_{p,q}$  is a mean. □

The proof of the following corollary follows immediately from Lemma 1.2.

**Corollary 3.1.** For  $0 < y < x$ , we have

$$(1) \quad \tilde{P}_{p,q}(x, y) < 2 \sin_{p,q} \left( \frac{x-y}{x+y} \right) \frac{A^2}{x-y}, \quad 2 \leq p < q,$$



$$(2) \quad \tilde{M}_{p,q}(x, y) < 2 \sinh_{p,q} \left( \frac{x-y}{x+y} \right) \frac{A^2}{x-y}, \quad 2 \leq p < q,$$

**Proof of Theorem 2.2.** Let  $f(t) = \sqrt{F(t)}$  and  $g(t) = \sqrt{G(t)}$  in Cauchy-Bouniakowski inequality (6), where  $F(t), G(t) > 0$ . Put  $[a, b] = [0, x]$ , then one gets the inequality:

$$(21) \quad \left( \int_0^x \sqrt{F(t)G(t)} dt \right)^2 \leq \int_0^x F(t) dt \cdot \int_0^x G(t) dt.$$

With the same notations, from the Pólya-Szegő inequality (7) one gets:

$$(22) \quad k(x, p) \left( \int_0^x \sqrt{F(t)G(t)} dt \right)^2 \geq \int_0^x F(t) dt \cdot \int_0^x G(t) dt,$$

here  $k(x, p)$  is as defined in Theorem 2.2. Let now  $f(t) = (1-t^p)^{-1/p}$  and  $g(t) = (1+t^p)^{-1/p}$ . From (21) and (22) one obtains

$$(23) \quad \arcsin_{2p}(x)^2 \leq \arcsin_p(x) \operatorname{arcsinh}_p(x),$$

and

$$(24) \quad \arcsin_p(x) \operatorname{arcsinh}_p(x) \leq k(x, p) \arcsin_{2p}(x)^2,$$

respectively. By definition, inequality (23) and (24) imply the proof of left hand-side and right-hand side of (12), respectively.  $\square$

**Proof of Theorem 2.3.** Apply the Diaz-Metcalf inequality (10) for  $f(t) = \sqrt{F(t)}$ ,  $g(t) = \sqrt{G(t)}$ ,  $[a, b] = [0, x]$ , yielding

$$\int_0^x G dt + M \cdot m \cdot \int_0^x F dt \leq (M + m) \cdot \int_0^x \sqrt{FG} dt.$$

Let  $F(t) = (1+t^q)^{-1/p}$ ,  $G(t) = (1-t^q)^{-1/p}$ . Here  $G(t)/F(t) = ((1+t^q)/(1-t^q))^{1/p}$ , which is strictly increasing. Thus

$$m = 1 \leq \sqrt{F/G} \leq ((1+x^q)/(1-x^q))^{1/(2p)} = M.$$

One obtains

$$(25) \quad \arcsin_p(x) + M \cdot \operatorname{arcsinh}_{p,q}(x) \leq (M+1) \arcsin_{2p,2q}(x),$$

this implies the proof of (13).

Let  $[a, b] = [0, x]$  and  $f(t) = (1+t^q)^{-1}$  and  $g(t) = (1-t^q)^{-1}$ . As  $f(t) + g(t) = 2/(1-t^{2q})$ , applying the Minkowski inequality (9) for  $t = 1/p$ ,  $p > 1$ , we get

$$(26) \quad \arcsin_{p,q}(x)^p + \operatorname{arcsinh}_{p,q}(x)^p \leq 2 \left( \int_0^x A^2 dt \right)^p,$$

where  $A(t) = 1/(1-t^{2q})^{1/(2p)}$ . Clearly,  $\int_0^x A(t) dt = \arcsin_{2p,2q}(x)$ , so for obtaining an upper bound for  $\int_0^x A(t)^2 dt$ , we apply the Pólya-Szegő inequality for  $f(t) = 1/(1-t^{2q})^{1/p}$  and  $g(t) = 1$ . Since in this case one has  $1 \leq f(t) \leq 1/(1-x^{2q})^{1/p}$ , we get from (7)

$$\int_0^x A(t)^2 dt \leq \arcsin_{2p,2q}(x)^2 R(x, p, q),$$

By using (26), finally we get

$$(27) \quad \frac{x^q (\arcsin_p(x)^p + \operatorname{arcsinh}_p(x)^p)}{\arcsin_{2p,2q}(x)^{2p}} \leq R(p, q, x),$$

this implies inequality (14).  $\square$

**Corollary 3.2.** For  $2 \leq p < q$ ,  $x = (a-b)/(a+b)$  and  $A = (a+b)/2$  with  $0 < b < a$ , we have

$$(28) \quad \frac{x A}{\min\{\frac{\pi_{p,q}}{2}, (1-x^q)^{-1/(p(1+q))}x\}} < \tilde{P}_{p,q} < \frac{A}{\left(1 + \frac{x^q}{p(1+q)}\right)}$$

$$(29) \quad \frac{(1+x^q)^{1/p} A}{s_{p,q}(x)} < \tilde{M}_{p,q} < \frac{(1+x^q)^{1/p} A}{r_{p,q}(x)},$$

where  $r_{p,q}(x)$  and  $s_{p,q}(x)$  are as in (4).

*Proof.* Proof follows easily from (28) and (29).  $\square$

4. CONCLUSION

In this short paper we generalized the Seiffert mean  $P$  and Neuman-Sándor mean  $M$  with two parameters in the form of Theorem 2.1, and proved some inequalities involving these means. We finish the paper by posing the following conjecture.

**Conjecture 4.1.** For  $p, q \geq 2$  and  $x = (a - b)/(a + b)$  with  $0 < b < a$ , the following functions

$$\tilde{L}_{p,q} = \tilde{L}_{p,q}(a, b) = \frac{a - b}{2\operatorname{arctanh}_{p,q}\left(\frac{a-b}{a+b}\right)} = \frac{x}{\operatorname{arctanh}_{p,q}(x)}A,$$

$$\tilde{T}_{p,q} = \tilde{T}_{p,q}(a, b) = \frac{a - b}{2\operatorname{arctan}_{p,q}\left(\frac{a-b}{a+b}\right)} = \frac{x}{\operatorname{arctan}_{p,q}(x)}A,$$

define a mean functions of two variables  $a$  and  $b$ , where  $\operatorname{arctanh}_{p,q}$  and  $\operatorname{arctan}_{p,q}$  are the inverse functions of  $\tanh_{p,q}$  and  $\tan_{p,q}$ , respectively.

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## Stancu type Dunkl generalization of Szász-Durrmeyer operators involving two variable hermite polynomials

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ABSTRACT. The aim of the present paper is to introduce a Stancu type Dunkl generalization of Szász- Durrmeyer operators involving two-variable Hermite polynomials defined by Krech [7]. Then, we give approximation properties for these operators with the help of Korovkin theorem. Furthermore, we obtain other approximation results via the class of Lipschitz functions, classical modulus of continuity, second order modulus of continuity and Peetre's K-functional.

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KEYWORDS: Dunkl exponential, Szász operators, modulus of continuity, Lipschitz functions, two-variable Hermite polynomials, Szász- Durrmeyer operators.

### 1. INTRODUCTION

In 1912, Bernstein [15] introduced the following sequence of operators  $B_n : B[0, 1] \rightarrow C[0, 1]$  defined by

$$(1) \quad B_n(f, x) = \sum_{k=0}^n \binom{n}{k} x^k (1-x)^{n-k} f\left(\frac{k}{n}\right),$$

where  $n \in \mathbb{N}$ ,  $x \in [0, 1]$  and  $f \in B[0, 1]$ . Some generalizations and modifications of the Bernstein polynomials can be found in [11], [9], [13]. Moreover, Szász [14] and Mirakjan [10] have introduced following operator

$$(2) \quad S_n(f, x) = e^{-nx} \sum_{k=0}^{\infty} \frac{(nx)^k}{k!} f\left(\frac{k}{n}\right),$$

where  $n \in \mathbb{N}$ ,  $x \geq 0$ ,  $f \in C[0, \infty)$ . For  $f \in C[0, \infty)$ , Durrmeyer type integral modifications of the operators (2) was defined by Mazhar and Totik [9] as follows

$$(3) \quad D_n(f, x) = n \sum_{j=0}^{\infty} e^{-nx} \frac{(nx)^j}{j!} \int_0^{\infty} e^{-nx} \frac{(nx)^k}{k!} f(t) dt,$$

where  $x \in [0, 1]$ . Also, Dunkl analogue of Szász operators given in [16] has been defined as

$$(4) \quad S_n^*(f, x) = \frac{1}{e_{\mu}(nx)} \sum_{k=0}^{\infty} \frac{(nx)^k}{\gamma_{\mu}(k)} f\left(\frac{k + 2\mu\theta_k}{n}\right),$$

where  $n \in \mathbb{N}$ ,  $x \geq 0$ ,  $f \in C[0, \infty)$ ,  $\mu \geq 0$ .

The Dunkl exponential function is given for  $\mu \geq -1/2$

$$(5) \quad e_{\mu}(nx) = \sum_{k=0}^{\infty} \frac{(nx)^k}{\gamma_{\mu}(k)},$$

where the coefficients are as follows

$$(6) \quad \gamma_\mu(2k) = \frac{2^{2k}k!\Gamma(k + \mu + \frac{1}{2})}{\Gamma(\mu + \frac{1}{2})} \quad \text{and} \quad \gamma_\mu(2k + 1) = \frac{2^{2k+1}k!\Gamma(k + \mu + \frac{3}{2})}{\Gamma(\mu + \frac{3}{2})}$$

and the recurrence relation

$$(7) \quad \frac{\gamma_\mu(k + 1)}{\gamma_\mu(k)} = (2\mu\theta_{k+1} + k + 1)$$

is satisfied where  $\theta_k$  is given by [13]

$$(8) \quad \theta_k = \begin{cases} 0 & \text{if } k = 2p \\ 1 & \text{if } k = 2p + 1 \end{cases}$$

for  $p \in N \cup \{0\}$ . The Dunkl generalization of two-variable Hermite polynomials [5] is defined by

$$(9) \quad \sum_{k=0}^{\infty} \frac{h_k^\mu(n, \alpha)}{\gamma_\mu(k)} x^k = e^{\alpha x^2} e_\mu(n, x),$$

where

$$(10) \quad h_k^\mu(n, \alpha) = \frac{\gamma_\mu(k) H_k^\mu(n, a)}{k!}.$$

Also, Wafi and Rao [18] defined the Szász-Gamma operators based on Dunkl analogue as follows

$$(11) \quad D_n f(x) = \frac{1}{e_\mu(n, x)} \sum_{k=0}^{\infty} \frac{(nx)^k}{\gamma_\mu(k)} \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k + 2\mu\theta_k + \lambda + 1)} \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nx} f(t) dt,$$

where  $x \in [0, \infty)$ ,  $\lambda \geq 0$  and the well-known Gamma function  $\Gamma(x)$  is defined as

$$(12) \quad \Gamma(x) = \int_0^\infty t^{x-1} e^{-t} dt.$$

Moreover, Dunkl-Gamma type operator in terms of generalization of two-variable Hermite polynomials [6] is defined by

$$(13) \quad S_n(f, x) = \frac{1}{e^{\alpha x^2} e_\mu(n, x)} \sum_{k=0}^{\infty} \frac{h_k^\mu(n, \alpha)}{\gamma_\mu(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k + 2\mu\theta_k + \lambda + 1)} \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} f(t) dt$$

where  $\alpha \geq 0$ ,  $\mu \geq -1/2$ ,  $\lambda \geq 0$  and  $x \in [0, \infty)$ .

In the present paper, we first give some lemmas and definitions to obtain convergence properties of the operators. Then, we define the Stancu type Dunkl generalization of Szász-Durrmeyer operators involving two-variable Hermite polynomials. Finally, we give the rates of convergence of the operator using the classical modulus of continuity, second order modulus of continuity, Peetre's K-functional and in terms of the elements of the Lipschitz class  $Lip_M(v)$ .

## 2. SOME LEMMAS AND DEFINITIONS

**Lemma 2.1.** *Let  $h_k^\mu(n, \alpha)$  be the Dunkl generalization of two-variable Hermite polynomials. Then the following equalities are hold.*



$$\begin{aligned}
 i) \quad & \sum_{k=0}^{\infty} \frac{h_{k+1}^{\mu}(n, \alpha)}{\gamma_{\mu}(k)} x^k = e^{\alpha x^2} e_{\mu}(nx)(2\alpha x + n) \\
 ii) \quad & \sum_{k=0}^{\infty} \frac{h_{k+2}^{\mu}(n, \alpha)}{\gamma_{\mu}(k)} x^k = e^{\alpha x^2} e_{\mu}(nx)(2\alpha^2 x^2 + 4\alpha n x + 2\alpha + n^2) + 4\alpha\mu e^{\alpha x^2} e_{\mu}(-nx) \\
 iii) \quad & \sum_{k=0}^{\infty} \frac{h_{k+3}^{\mu}(n, \alpha)}{\gamma_{\mu}(k)} x^k = e^{\alpha x^2} e_{\mu}(-nx)(8\alpha^2 x\mu + 4\alpha n\mu) + e^{\alpha x^2} e_{\mu}(nx) + e^{\alpha x^2} e_{\mu}(nx)(8\alpha^3 x^3 \\
 & + 12\alpha^2 x^2 n + 12\alpha^2 x + 6\alpha n x^2 + 6\alpha n + n^3).
 \end{aligned}$$

*Proof.* The proofs of above equalities are given in [3, 17]. □

**Definition 2.2.** With the help of the Dunkl generalization of two variable Hermite polynomials given (9), we introduce the operators  $D_n^H(f, x)$

$$(14) \quad D_n^H(f, x) = \frac{1}{e^{\alpha x^2} e_{\mu}(nx)} \sum_{k=0}^{\infty} \frac{h_k^{\mu}(n, \alpha)}{\gamma_{\mu}(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \int_0^{\infty} t^{k+2\mu\theta_k+\lambda} e^{-nt} f\left(\frac{nt+\gamma}{n+\beta}\right) dt,$$

where  $x \in [0, \infty)$ ,  $\alpha \geq 0, \mu > -1/2$ ,  $\lambda \geq 0$  and  $\gamma, \beta$  parameters satisfy the condition  $0 \leq \gamma \leq \beta$ . We note that as a special case  $\gamma = \beta = 0$ , then  $D_n^H$  reduces to  $S_n$  operator defined (13).

**Lemma 2.3.** For the positive linear operator  $D_n^H(f, x)$  given by (14), then the following results are satisfied.

$$\begin{aligned}
 i) \quad & D_n^H(1, x) = 1 \\
 ii) \quad & D_n^H(t, x) = \frac{nx}{n+\beta} + \frac{2\alpha x^2}{n+\beta} + \frac{\lambda+1}{n+\beta} + \frac{\gamma}{n+\beta} \\
 iii) \quad & D_n^H(t^2, x) = \frac{4\alpha^2 x^4}{(n+\beta)^2} + \frac{4n\alpha x^3}{(n+\beta)^2} + \frac{(10\alpha+n^2+4\lambda\alpha+4\gamma\alpha)x^2}{(n+\beta)^2} + \frac{(2\lambda n+4n+2\gamma n)}{(n+\beta)^2} \\
 & + \left(\frac{2\mu n-4\alpha\mu x^2+4\alpha\mu}{(n+\beta)^2}\right) \frac{e_{\mu}(-nx)}{e_{\mu}(nx)} + \frac{\lambda^2+3\lambda+2}{(n+\beta)^2} + \frac{2\gamma\lambda+2\gamma}{(n+\beta)^2}.
 \end{aligned}$$

*Proof.* **i)** By using the generating function (9), the relation (i) holds.

**ii)** With the help of definition of the Dunkl analogue of two-variable Hermite polynomials and the Gamma function, we get

$$\begin{aligned}
 D_n^H(f, x) &= \frac{n}{(n+\beta)e^{\alpha x^2} e_{\mu}(nx)} \sum_{k=0}^{\infty} \frac{h_k^{\mu}(n, \alpha)}{\gamma_{\mu}(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \frac{\Gamma(k+2\mu\theta_k+\lambda+2)}{n^{k+2\mu\theta_k+\lambda+1}} \\
 &+ \frac{\gamma}{(n+\beta)} \sum_{k=0}^{\infty} \frac{h_k^{\mu}(n, \alpha)}{\gamma_{\mu}(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \frac{\Gamma(k+2\mu\theta_k+\lambda+1)}{n^{k+2\mu\theta_k+\lambda+1}}.
 \end{aligned}$$

Also by using recurrence relation (7), we have

$$D_n^H(f, x) = \frac{n}{(n+\beta)} + \frac{2\alpha x^2}{(n+\beta)} + \frac{\lambda+1}{n+\beta} + \frac{\gamma}{n+\beta}$$

iii) Similarly for  $f(t) = t^2$ ,

$$\begin{aligned}
 D_n^H(f, x) &= \frac{1}{e^{\alpha x^2} e_{\mu}(nx)} \sum_{k=0}^{\infty} \frac{h_k^{\mu}(n, \alpha)}{\gamma_{\mu}(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \\
 &\times \left( n^2 \int_0^{\infty} t^{k+2\mu\theta_k+\lambda+2} e^{-nt} dt + 2n\gamma \int_0^{\infty} t^{k+2\mu\theta_k+\lambda+1} e^{-nt} dt + \gamma^2 \int_0^{\infty} t^{k+2\mu\theta_k+\lambda+2} e^{-nt} dt \right),
 \end{aligned}$$

where  $\theta_{(k+1)} = \theta_k + (-1)^k$  and using the recurrence relation (7) completes the proof. □

**Theorem 2.4.** The operator  $D_n^H$  and any uniformly continuous bounded function  $g$  on the interval  $[0, \infty)$ , we can give

$$D_n^H \rightrightarrows g(x)$$

on each compact subset of  $[0, \infty)$  when  $n \rightarrow \infty$ .

*Proof.* From the results obtained in Lemma 2

$$\lim_{n \rightarrow \infty} D_n^H(e^i; x) = x^i \quad i = 0, 1, 2$$

holds. This convergence is uniformly in each compact subset of  $[0, \infty)$ . Then, applying the universal Korovkin type theorem gives the desired result.  $\square$

**Lemma 2.5.** *The operators  $D_n^H(t; x)$  satisfy the following results from Lemma 2.*

$$\begin{aligned} \Delta_1 &= D_n^H((t-x); x) = \frac{nx}{(n+\beta)} + \frac{2\alpha x^2}{(n+\beta)^2} + \frac{(\lambda+1)}{(n+\beta)} + \frac{y}{(n+\beta)} - x, \\ \Delta_2 &= D_n^H((t-x)^2; x) = \frac{4\alpha^2}{(n+\beta)^2} x^4 + \frac{4\alpha n}{(n+\beta)} x^3 + \frac{(10\alpha + n^2 + 4\lambda\alpha + 4\gamma\alpha)}{(n+\beta)^2} x^2 \\ &\quad + \frac{(2\lambda n + 4n + 2\gamma n)}{(n+\beta)^2} x + \frac{(2\mu n x - 4\alpha\mu x^2 + 4\alpha\mu)}{(n+\beta)^2} \frac{e_\mu(-nx)}{e_\mu(nx)} + \frac{\lambda^2 + 3\lambda + 2}{(n+\beta)^2} + \frac{2\gamma\lambda + 2\gamma}{(n+\beta)^2} \\ &\quad - 2x \left( \frac{nx + 2\alpha x^2 + (\lambda+1) + \gamma}{(n+\beta)} \right) + x^2. \end{aligned}$$

**Lemma 2.6.** *The following inequality holds true for  $h \in C_\infty^2[0, \infty]$*

$$D_n^H(h; x) - h(x) \leq [\Delta_1 + \Delta_2] \|h\|_{C_\infty^2},$$

where  $\Delta_1$  and  $\Delta_2$  are central moments of operator  $D_n^H$  given by in Lemma 3.

*Proof.*  $\square$

Lagrange form of the remaining piece of the Taylor series

$$h(s) = h(x) + (s-x)h'(x) + \frac{(s-x)^2}{2!} h''(\sigma), \quad \sigma \in (x, s).$$

Applying the operator  $D_n^H$  to both sides of this inequality and then using the linearity of the operator, we have

$$D_n^H(h; x) - h(x) = h'(x)\Delta_1 + \frac{h''(\sigma)}{2}\Delta_2$$

and, it yields

$$|D_n^H(h; x) - h(x)| \leq \|h'(x)\| \Delta_1 + \|h''(x)\| \Delta_2 \leq [\Delta_1 + \Delta_2] \|h\|.$$

### 3. SPECIAL CONVERGENCE RESULTS

In this part, we give some definitions and rates of convergence results of the operators  $D_n^H(f; x)$ .

Let  $g \in Lip_M(\nu)$  then the following inequality is hold

$$|g(s) - g(t)| \leq M |s - t|^\nu,$$

where  $s, t \in [0, \infty)$ ,  $0 < \nu \leq 1$ ,  $M > 0$ .

The modulus of continuity is defined by

$$w(g; \delta) := \sup_{\substack{s, t \in [0, \infty) \\ \|s-t\| \leq \delta}} |g(s) - g(t)|,$$

where  $g \in \check{C}[0, \infty)$  which is the space of uniformly continuous on  $[0, \infty)$ .

**Theorem 3.1.** *If  $f \in Lip_M(\nu)$ , then we have*

$$|D_n^H(f; x) - f(x)| \leq M(\Delta_2)^{\nu/2},$$

where  $\Delta_2 = D_n^H((t-x)^2; x)$ .

*Proof.* With the help of the definition of  $f \in Lip_M(v)$  and linearity of operator  $D_n^H$ , we get

$$\begin{aligned} D_n^H(f(t) - f(x); x) &\leq MD_n^H(|t - x|^v; x) \\ D_n^H(f(t); x) - f(x) &\leq MD_n^H(|t - x|^v; x) \\ |D_n^H(f(t); x) - f(x)| &\leq MD_n^H(|t - x|^v; x). \end{aligned}$$

Let use the definition of operator

$$\begin{aligned} D_n^H(|t - x|; x) = D_n^H(f, x) &= \frac{1}{e^{\alpha x^2} e_\mu(nx)} \sum_{k=0}^{\infty} \frac{h_k^\mu(n, \alpha)}{\gamma_\mu(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \\ &\times \left( \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} \left( \frac{nt+\gamma}{n+\beta} - x \right)^v dt \right) \end{aligned}$$

and using Hölder's inequality for the above integral

$$\begin{aligned} &\leq \frac{1}{e^{\alpha x^2} e_\mu(nx)} \sum_{k=0}^{\infty} \frac{h_k^\mu(n, \alpha)}{\gamma_\mu(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \\ &\times \left( \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} \left( \frac{nt+\gamma}{n+\beta} - x \right)^2 dt \right)^{v/2} \left( \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} \right)^{\frac{2-v}{2}}. \end{aligned}$$

and then, we apply Hölder's inequality for sum

$$\begin{aligned} &\leq \left( \sum_{k=0}^{\infty} \frac{1}{e^{\alpha x^2} e_\mu(nx)} \frac{h_k^\mu(n, \alpha)}{\gamma_\mu(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} \left( \frac{nt+\gamma}{n+\beta} - x \right)^2 dt \right)^{v/2} \\ &\times \left( \sum_{k=0}^{\infty} \frac{1}{e^{\alpha x^2} e_\mu(nx)} \frac{h_k^\mu(n, \alpha)}{\gamma_\mu(k)} x^k \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} \right)^{\frac{2-v}{2}} \end{aligned}$$

by using the above inequalities, we get

$$D_n^H(|t - x|^v; x) \leq D_n^H(|t - x|^2; x)^{\frac{v}{2}}$$

from which, it follows

$$|D_n^H(f(t); x) - f(x)| \leq MD_n^H(|t - x|^v; x) \leq MD_n^H(|t - x|^2; x)^{\frac{v}{2}}$$

$$|D_n^H(f(t); x) - f(x)| \leq MMD_n^H(|t - x|^2; x)^{\frac{v}{2}}$$

□

**Theorem 3.2.**  $D_n^H((g(t); x)$  operators satisfy the inequality

$$|D_n^H(g(t); x) - g(x)| \leq 2w(g; \sqrt{\Delta_2}),$$

where  $g \in C_B[0, \infty)$ .

*Proof.*

□

As a consequence of classical modulus of continuity and  $g \in C_B[0, \infty)$

$$|g(t) - g(x)| \leq w(g; \delta) \left( \frac{|t - x|}{\delta} + 1; x \right)$$

since  $D_n^H$  operator is also positive linear operator, we can write

$$\begin{aligned} |D_n^H(g; x) - g(x)| &= |D_n^H(g(t) - g(x); x)| \leq D_n^H(|g(t) - g(x)|; x) \leq w(g; \delta) D_n^H\left(\frac{t - x}{\delta}; x\right) \\ &= w(g; \delta) \left( \frac{D_n^H(|t - x|; x)}{\delta} + 1 \right). \end{aligned}$$

In addition to this inequality, we use the definition of operator

$$D_n^H(|t-x|; x) = \frac{1}{e^{\alpha x^2} e_\mu(nx)} \sum_{k=0}^{\infty} \frac{h_k^\mu(n, \alpha)}{\gamma_\mu(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \\ \times \left( \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} \left( \frac{nt+\gamma}{n+\beta} - x \right) dt \right)$$

then by applying Cauchy Schwarz inequality

$$D_n^H(|t-x|; x) \leq \frac{1}{e^{\alpha x^2} e_\mu(nx)} \sum_{k=0}^{\infty} \frac{h_k^\mu(n, \alpha)}{\gamma_\mu(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \\ \times \left( \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} \left( \frac{nt+\gamma}{n+\beta} - x \right)^2 dt \right)^{\frac{1}{2}} \left( \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} dt \right)^{\frac{1}{2}}.$$

and using Cauchy- Schwarz inequality for summation on right hand side

$$\leq \left( \sum_{k=0}^{\infty} \frac{1}{e^{\alpha x^2} e_\mu(nx)} \frac{h_k^\mu(n, \alpha)}{\gamma_\mu(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} \left( \frac{nt+\gamma}{n+\beta} - x \right)^2 dt \right) \\ \times \left( \sum_{k=0}^{\infty} \frac{1}{e^{\alpha x^2} e_\mu(nx)} \frac{h_k^\mu(n, \alpha)}{\gamma_\mu(k)} x^k \frac{n^{k+2\mu\theta_k+\lambda+1}}{\Gamma(k+2\mu\theta_k+\lambda+1)} \int_0^\infty t^{k+2\mu\theta_k+\lambda} e^{-nt} dt \right)^{\frac{1}{2}}$$

$$D_n^H(|t-x|; x) \leq \sqrt{\Delta_2}.$$

Thus, we can give the following inequality

$$w(g; \delta) \left( \frac{D_n^H(|t-x|; x)}{\delta} + 1 \right) \leq w(g; \delta) \left( \frac{\sqrt{\Delta_2}}{\delta} + 1 \right)$$

so,

$$|D_n^H(g; x) - g(x)| \leq 2w(g; \delta),$$

where  $\delta = \sqrt{\Delta_2}$ . Hence, the proof complete.

**Theorem 3.3.** For the operators  $D_n^H(f; x)$ , the following inequality holds

$$|D_n^H(f; x) - g(x)| \leq 2K \left( g; \left( \sqrt{\Delta_1} + \frac{\sqrt{\Delta_2}}{2} \right) \right),$$

where  $\Delta_1, \Delta_2$  are central moments of  $D_n^H$ .

*Proof.* Lagrange form of the remaining piece of the Taylor series for  $f \in C_B^2[0, \infty)$

$$f(t) - f(x) = f'(x)(t-x) + f''(c) \frac{(t-x)^2}{2} \quad x \leq c \leq t$$

Then, by applying the operator  $D_n^H(f; x)$  both side

$$|D_n^H(f; x) - g(x)| = |f'(x)| D_n^H(|t-x|; x) + \frac{|f''(c)|}{2} D_n^H(t-x)^2; x)$$

since  $D_n^H(|t-x|; x) \leq \sqrt{\Delta_2}$ , we get

$$|D_n^H(f; x) - g(x)| \leq |f'(x)| \sqrt{\Delta_1} + \frac{|f''(c)|}{2} \sqrt{\Delta_2} \\ |D_n^H(f; x) - g(x)| \leq \|f'(x)\|_{C_B[0, \infty)} \sqrt{\Delta_1} + \frac{\|f''(c)\|_{C_B[0, \infty)}}{2} \sqrt{\Delta_2} \\ |D_n^H(f; x) - g(x)| \leq \left( \sqrt{\Delta_1} + \frac{\sqrt{\Delta_2}}{2} \right) \|f\|_{C_B^2[0, \infty)}.$$

By using the above inequality, we can write

$$|D_n^H(f; x) - g(x)| = |D_n^H(g; x) - D_n^H(f; x) + D_n^H(f; x) - f(x) + f(x) - g(x)| \\ \leq D_n^H(|g-f|; x) + |g(x) - f(x)| + |D_n^H(f; x) - f(x)|.$$

$$(15) \quad |D_n^H(f; x) - g(x)| \leq 2 \|g - f\|_{C_B[0, \infty)} + 2 \|f\|_{C_B^2[0, \infty)} \left( \sqrt{\Delta_1} + \frac{\sqrt{\Delta_2}}{2} \right).$$

The definition of Petre's  $K$  functional of the function  $f \in C_B [0, \infty)$  is given as follows

$$K(f; \delta) := \inf_{g \in C_B^2[0, \infty)} \left\{ \|f - g\|_{C_B[0, \infty)} + \delta \|g\|_{C_B[0, \infty)} \right\},$$

where for all  $\delta > 0$ . By taking infimum of both sides of (15) of for  $f$ , it yields

$$|D_n^H(f; x) - g(x)| \leq 2K \left( g; \sqrt{\Delta_1} + \frac{\sqrt{\Delta_2}}{2} \right)$$

Hence, the proof complete. □

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# A mathematical model for the effects of wavelets and the analysis of neural network operators described using wavelets

Harun Karsh

ABSTRACT. It is very well-known that wavelets have great advantage of being able to separate and identify fine details in a signal or a function. One of the main advantages of wavelets compared to the Fourier analysis and its related theories is that they offer simultaneous localization in the time and frequency domain. The second main advantage of wavelets is that they are computationally very fast and detailed when using wavelet expansions and transformations. In the present study we will deal with the linear approximation operators constructed by compactly supported Daubechies wavelets. In details, we will reconstruct neural network operators, where location and time are very important and effective, with the help of wavelets, and we will examine and analyse various properties of the wavelet type extension of the neural network operators.

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## 1. INTRODUCTION

It is known that Artificial Neural Networks (ANN) are defined as a mathematical model that allows the brain's behavior and abilities to be reproduced. In general, Neural Network is a family of parametrized functions, namely by a mathematical point of view it is a multiple composition of some special functions called activations. It is important to point out that, this approach motivated by the fundamental density theorem due to George Cybenko given in [23].

In 1989, Cybenko [23] gave an answer to the superposition problem on  $C[a, b]$  with his famous density theorem, which states that every continuous function defined on  $[a, b]$  can be approximated by a sequence constructed by a linear combination of sigmoidal functions. In other words, Cybenko confirmed that a neural network with solely one hidden-layer is capable of always approximating a continuous function. The main advantage of the above theory lies on its connection with the Artificial Neural Networks, Learning Theory and their applications to Approximation Theory, see e.g. [4], [22] and [33].

Based on the idea developed by Cybenko, the theory of the mathematical models of the neural network (NN) operators arise since 1992 with the pioneer work of Cardaliaguet and Euvrard [12], and then in the next years, they have been largely studied by several authors under different aspects. Especially in the last two decades, there are many new version of artificial neural networks has been introduced and widely studied.

In particular, in 1997 Anastassiou [2] pointed out and obtained that the compactly supported bell-shaped functions used in the Density Theorem of Cybenko and the Cardaliaguet and Euvrard (NN) operators can be obtained from sigmoidal functions used effectively in Artificial Neural Networks, serious relations have emerged between the Cybenko convergence theorem and the Theory of Approximations. Thanks to this fundamental work of Anastassiou in 1997, the techniques and theorems of the Theory of Approximations began to be used in Artificial Neural Networks.

Owing to the work of the famous mathematician G. Anastassiou in 1997, see [2], the theory of neural network(NN) operators has been introduced by Anastassiou as an extension of the bell-shaped operators studied by Cardaliaguet and Euvrard in [12].

Especially in the last two decades, based on the idea developed by Anastassiou [2] and afterwards the continuous works on these operators and some of their modifications of D.

Costarelli, R. Spigler, C. Bardaro, G. Vinti and their research group (RITA network) from Perugia, such as Kantorovich and Durrmeyer forms have been of great importance in the development of mathematical models for signal and image recovering. And hence the approximation problem was proved by using NN operators and some of its different forms in some function spaces, see e.g., [3]-[6], [10]-[11], [14]-[21], [7], [9] and [13]). Since the classical neural network operators cannot be used for  $L^p(\mathbb{R})(1 \leq p < \infty)$  approximation, to obtain some positive results for these functions, their Kantorovich and Durrmeyer type modifications are considered.

Especially, in the paper [15], Costarelli and Spigler introduced and investigated a kind of neural network operators for absolutely continuous functions, by considering the relations between sigmoidal functions, multiresolution analysis and the scaling functions. Moreover they also state some interesting results related with the rate of convergence in the set  $AC[a, b]$ .

The goal of this study is to find a positive solution to the approximation (or superposition) problem for operators in some general function spaces, by using the effects and relations between different function spaces provided by the wavelets. In other words, we will propose a generalization and extension of the theory of interpolation to operators by introducing an integral operator, called wavelet type operators (see [1], [26], [31], [32] and [35]). The new operators are more flexible than the previous ones and they are at least a natural extensions of the classical NN operators, Kantorovich and Durrmeyer type modifications.

The basis used in this construction of the new neural network operators are the Cybenko density theorem and wavelets.

Similar constructions and investigations can be found in the very recent papers of the author [27]-[30].

It is very well-known that wavelets and wavelet expansions have the great advantage of being able to separate and identify fine details in a signal or a function.

One of the main advantages of wavelets compared to the Fourier analysis and its related theories is that they offer simultaneous localization in the time and frequency domain. The second main advantage of wavelets is that they are computationally very fast and detailed when using wavelet expansions and transformations.

Unlike the Fourier analysis, wavelets tell us about the frequencies present as well as the time in which these frequencies were observed.

So, wavelets are a better way of analyzing especially the dynamic signals because they have a relatively higher resolution in both time and frequency domain.

Moreover, from the definitions and properties of the wavelet bases, one can use wavelet type operators for approximation problem in  $L^p$  spaces.

Since wavelets have many advantages for approximating in  $L^p$  spaces, potential applications in machine learning and neural networks, the future directions of this work are to try to adapt what has been accomplished in wavelets to these spaces and theories (see [5] and [22]).

In Sect. 2, we recall the definition of the NN operators and wavelets, and all their main properties which are useful in order to prove the quantitative estimates, together with some examples of activation functions, while in Sect. 3 the main results of the paper have been established. In the final part of this study, we will provide also some graphical examples and comparisons between the convergence of the NN operators obtained different kind of wavelets.

## 2. PRELIMINARIES AND AUXILIARY RESULTS

In this section we shall recall some notation and background material of the theory of Neural Networks and the theory of wavelet, especially Daubechies' compactly supported wavelets ([24], [25]), used throughout this paper.

We denote by  $C[a, b]$ ,  $B[a, b]$  and  $L_\infty(\mathbb{R})$  the sets of continuous, bounded and essentially bounded functions with their usual norms, respectively.

**Definition 2.1. (Centered bell-shaped function)** A function  $b : \mathbb{R} \rightarrow \mathbb{R}$  is said to be centered bell-shaped if  $b$  belongs to  $L^1$  and its integral is nonzero, if it is nondecreasing on  $(-\infty, 0)$  and nonincreasing on  $[0, +\infty)$ .

**Definition 2.2. (Sigmoidal Function)** Let  $\sigma : \mathbb{R} \rightarrow \mathbb{R}$  be a measurable function satisfies

$$\lim_{x \rightarrow -\infty} \sigma(x) = 0 \quad \text{and} \quad \lim_{x \rightarrow \infty} \sigma(x) = 1,$$

then it is called sigmoidal function.

As an example of a sigmoidal and its corresponding bell-shaped function, we can consider the ramp function, which is very useful and important for the neural network.

**Definition 2.3. (Ramp Function)** A ramp function is a special sigmoidal function defined as

$$R(x) = \begin{cases} 0 & , \quad x \leq -1/2 \\ x + 1/2 & , \quad -1/2 < x < 1/2 \\ 1 & , \quad x \geq 1/2 \end{cases} .$$

Clearly, a bell-shaped function can be define by using Sigmoidal ( or Ramp) function, namely

$$b_\sigma(x) = \frac{\sigma(x+1) - \sigma(x-1)}{2},$$

and

$$b_R(x) = R(x+1/2) - R(x-1/2).$$

Moreover,  $b_R : \mathbb{R} \rightarrow \mathbb{R}$  is a bell-shaped kernel function obtained by ramp function  $R(x)$  that satisfies following assumptions:

$b_R$  is a continuous function on  $\mathbb{R}$ ,

$$(1) \quad b_R \in L^1(\mathbb{R}), \quad \sum_{k=0}^n b_R(u-k) = 1 \text{ for every } u \in \mathbb{R},$$

and

$$(2) \quad O_R := \sup_{u \in \mathbb{R}} \sum_{k=0}^n b_R(u-k) < \infty$$

where the convergence of the series (2) is uniform on each compact subintervals of  $\mathbb{R}$ .

Some other sigmoidal functions are Gompertz function, Logistic function, Error function and Hyperbolic tangent functions, etc.

It is important to note that, the theory of sigmoidal functions are not new. At 1838, the logistic function was introduced by Pierre Franois Verhulst [36]-[37], who applied it to human population dynamics. Verhulst derived his logistic equation to describe the mechanism of the self-limiting growth of a biological population.

Since then the logistic functions have many applications in many research areas, including biology, ecology, population dynamics, chemistry, demography, economics, geoscience, mathematical psychology, probability, sociology, political science, nancial mathematics, statistics.

Now we will give some definitions of the Neural Network (NN) Operators.

**Definition 2.4. (Cardaliaguet and Euvrard (NN) Operators)**

Let  $f : \mathbb{R} \rightarrow \mathbb{R}$  be a continuous and bounded function and  $n \in \mathbb{N}^+$ , the Cardaliaguet and Euvrard (NN) Operators are defined as.

$$(F_n f)(x) = \sum_{k=-n^2}^{n^2} \frac{f\left(\frac{k}{n}\right)}{B n^\alpha} b(n^{-\alpha}(nx - k)),$$

where  $0 < \alpha < 1$ ,  $b$  is a bell-shaped function with compact support  $\subset [-T, T]$ , and

$$B := \int_{-T}^T b(x) dx$$

The Cardaliaguet and Euvrard (NN) Operators and Its different modifications were intensively studied by Anastassiou, Spiegler, Costarelli, Vinti.

In [2], Anastassiou defined the Cardaliaguet and Euvrard (NN) Operators as follows;

**Definition 2.5.** Let  $f : \mathbb{R} \rightarrow \mathbb{R}$  be a continuous and bounded function,  $T > 0, n \in \mathbb{N}^+$  and  $n \geq \max\{T + |x|, T^{-1/\alpha}\}$ , the Cardaliaguet and Euvrard (NN) Operators are given by

$$(F_n f)(x) = \sum_{k=\lceil nx-Tn^\alpha \rceil}^{\lfloor nx+Tn^\alpha \rfloor} \frac{f\left(\frac{k}{n}\right)}{Bn^\alpha} b(n^{-\alpha}(nx-k)),$$

where again  $0 < \alpha < 1$ ,  $b$  is a bell-shaped function with compact support  $\subset [-T, T]$ , and

$$B := \int_{-T}^T b(x) dx.$$

As a special case of the operators defined on  $[a, b]$  and  $[0, 1]$ , we have the following type Neural Network (NN) Operators, respectively.

**Definition 2.6. (Neural Network (NN) Operators).** Let  $f : [a, b] \rightarrow \mathbb{R}$  be a bounded function, and  $n \in \mathbb{N}^+$  such that  $\lceil na \rceil \leq \lfloor nb \rfloor$ . The positive linear neural network operators activated by the ramp function  $R(x)$ , are defined as.

$$(F_n f)(x) = \frac{\sum_{k=0}^n f\left(a + k \frac{b-a}{n}\right) b_R\left(\frac{n(x-a-k \frac{b-a}{n})}{b-a}\right)}{\sum_{k=0}^n b_R\left(\frac{n(x-a-k \frac{b-a}{n})}{b-a}\right)}, \quad x \in [a, b],$$

and

**Definition 2.7. (Neural Network (NN) Operators).** Let  $f : [0, 1] \rightarrow \mathbb{R}$  be a bounded function, and  $n \in \mathbb{N}^+$ . The positive linear neural network operators activated by the ramp function  $R$ , are defined as.

$$(3) \quad (N_n f)(x) = \frac{\sum_{k=0}^n f\left(\frac{k}{n}\right) b_R(nx-k)}{\sum_{k=0}^n b_R(nx-k)}, \quad x \in [0, 1],$$

where  $b_R$  is the bell-shaped function obtained by the ramp function  $R$ .

Now, we will give some informations about wavelets, multiresolution analysis and its theories, which will be useful for the remain part of this work.

**Definition 2.8. (Scale Function)** A scale function defined on the interval  $[0, 1]$  as

$$\phi(t) = \begin{cases} 1 & , \quad 0 \leq x < 1 \\ 0 & , \quad e.w. \end{cases}.$$

Clearly, a scale function can be also define by using Heaviside unit step function, namely

$$H(x) = \begin{cases} 1 & , \quad x \geq 0 \\ 0 & , \quad x < 0 \end{cases},$$

and hence

$$\phi(x) = H(x) - H(x-1).$$

A multiresolution analysis (MRA) is a sequence  $(V_j)_{j \in \mathbb{Z}}$  of closed subspaces of  $L^2(\mathbb{R})$  such that the following hold:

i)  $V_j$  is a set of all  $f \in L^2(\mathbb{R})$  which are constant on  $2^{-j}$  length intervals and

$$\dots \subset V_{-2} \subset V_{-1} \subset V_0 \subset V_1 \subset \dots \subset V_j \subset V_{j+1} \subset \dots \subset L_2(\mathbb{R}),$$

$$\overline{\bigcup_j V_j} = L_2(\mathbb{R}),$$

ii)

$$\begin{aligned}\forall j, k \in \mathbb{Z}, f(x) \in V_j &\Leftrightarrow f(2x) \in V_{j+1}, \\ \forall k \in \mathbb{Z}, f(x) \in V_0 &\Leftrightarrow f(x - k) \in V_0 \\ \forall j, k \in \mathbb{Z}, f(x) \in V_j &\Leftrightarrow f(x - 2^{-j}k) \in V_j,\end{aligned}$$

and

iii)

$$\bigcap_j V_j = \{0\}.$$

Since  $V_j \subset V_{j+1}$  ( $\forall j \in \mathbb{Z}$ ) then there exist subspaces  $W_j$  of  $L_2(\mathbb{R})$  satisfying

$$\begin{aligned}V_1 &= V_0 \oplus W_0, \\ V_2 &= V_1 \oplus W_1,\end{aligned}$$

...

$$V_{j+1} = V_j \oplus W_j,$$

and

$$L_2(\mathbb{R}) = \dots W_{-2} \oplus W_{-1} \oplus W_0 \oplus \dots \oplus W_j \oplus W_{j+1} \oplus \dots$$

**Definition 2.9. (Wavelet)** A wavelet is a small wave which oscillates and decays in the time domain. A wavelet basis set starts with two orthogonal functions: the scaling function or father wavelet  $\phi(t)$  and the wavelet function or mother wavelet  $\vartheta(t)$ . By scaling and translation of these two orthogonal functions we obtain a complete basis set. The scaling and wavelet functions, respectively, satisfy

$$\int_{-\infty}^{\infty} \phi(t) dt = 1, \quad \int_{-\infty}^{\infty} \vartheta(t) dt = 0.$$

These two functions have finite energy, namely  $\phi, \vartheta \in L^2(\mathbb{R})$ , and orthogonal.

In general, the wavelets refers to the set of family of orthonormal functions of the form

$$\begin{aligned}\phi_{a,b}(t) &= \frac{1}{\sqrt{a}} \phi\left(\frac{t-b}{a}\right), \quad a > 0, b \in \mathbb{R}, \\ \vartheta_{a,b}(t) &= \frac{1}{\sqrt{a}} \vartheta\left(\frac{t-b}{a}\right), \quad a > 0, b \in \mathbb{R},\end{aligned}$$

where  $\phi$  and  $\vartheta$  are the basic, father and mother wavelets, respectively.

**Haar Wavelet:** The simplest wavelet is known as the Haar Wavelet defined as;

$$\vartheta(x) = \begin{cases} 1 & , \quad 0 \leq x < \frac{1}{2} \\ -1 & , \quad \frac{1}{2} \leq x < 1 \\ 0 & , \quad e.w. \end{cases}$$

with the scaling function

$$\phi(t) = \begin{cases} 1 & , \quad 0 \leq x < 1 \\ 0 & , \quad e.w. \end{cases}.$$

Clearly, Haar wavelets constitutes an orthonormal system for the space of square-integrable functions on the real line. Since Haar wavelet is not continuous and therefore not differentiable, it is suitable for representing discrete signals not for representing smooth signals or functions.

In the present study, we consider orthonormal bases of wavelets in  $L^2(\mathbb{R})$ , and assume that there is a scaling function (father wavelet)  $\phi(t)$  whose whose translates  $\{\phi(t - n)\}$  are orthogonal and the mother wavelet  $\vartheta(t)$  based on the father wavelet  $\phi(t)$  gives rise to the orthonormal basis  $\vartheta_{j,k}(t)$  of  $L^2(\mathbb{R})$ , where

$$(4) \quad \vartheta_{j,k}(t) = 2^{j/2} \vartheta(2^j t - k).$$



Hence, by using a multiresolution analysis (MRA), each  $f \in L^2(\mathbb{R})$  has the following representation

$$f(x) = \sum_{j \in \mathbb{Z}} \sum_{k \in \mathbb{Z}} b_{j,k} \vartheta_{j,k}(x),$$

called wavelet expansion of  $f \in L^2(\mathbb{R})$ , where  $b_{j,k}$  are wavelet coefficients defined by

$$b_{j,k} = \langle f(x), \vartheta_{j,k}(x) \rangle = 2^{j/2} \int_{\mathbb{R}} f(x) \overline{\vartheta(2^j x - k)} dx.$$

Some convergence results about wavelet expansions, please see [31], [35] and [32].

Some of the special cases of  $a$  and  $b$ , one can obtain different type of wavelets.

**Franklin system:** Let  $k, j \in \mathbb{Z}$ . If one choose  $a = 2^{-j}$  and  $b = k2^{-j}$ , then one can obtain the Franklin system, which is an orthonormal basis of  $L^2[0, 1]$ .

**Strömberg wavelet:** Even though the Haar wavelet was earlier known to be an orthonormal wavelet, Strömberg wavelet was the first smooth orthonormal wavelet to be discovered. Namely, J.O. Strömberg constructed the orthonormal basis of the form (4) by using a function  $\vartheta \in C^m$  for an arbitrary nonnegative integer  $m$ , which is a complete orthonormal system in the space of square integrable functions over  $\mathbb{R}$ .

The wavelet analysis procedure is to adopt a wavelet prototype function, called an analyzing wavelet (father wavelet) or mother wavelet.

**Definition 2.10.** ([24], [25]) (*Compactly supported Daubechies Wavelet*)

Owing to the above definitions, first of all, we introduce the NN operators by using the compactly supported Daubechies wavelets considered in this paper.

Let us assume that the scale function (or father wavelets)  $\psi \in L^\infty(\mathbb{R})$  and satisfies:

- a)  $\psi$  is a compactly supported, namely there is a real constant  $\lambda > 0$  such that  $\text{supp } \psi \subset [0, \lambda]$ ,
- b)  $\int_{-\infty}^{\infty} \psi(x) dx = 1$ ,
- c) the first  $N$  moments of the father wavelet  $\psi$  satisfy

$$\int_{-\infty}^{\infty} x^j \psi(x) dx = 0, \quad j = 1, \dots, N.$$

**Note 1.** Actually, Daubechies wavelets have strong relations with the properties of continuity and differentiability. Namely, for an arbitrary fixed integer  $N \geq 1$ , compactly supported Daubechies wavelet  $\psi$  is supported with  $[0, 2N - 1]$ , in addition there exists a constant  $r > 0$  such that for  $N \geq 2$ ,  $\psi \in C^{rN}(\mathbb{R})$  and to have a given number of vanishing moments.

In particular, when  $N = 1$ , then the first Daubechies wavelet  $\psi$  will be the classical Haar basis. As  $N$  increases, the regularity of the wavelets increase (see [24], [25]).

This means that if we want to use Daubechies wavelets to reconstruct a function, it is more convenient to choose or construct wavelets based on the continuity or differentiability properties of the given function (please see Haar wavelet).

Now, we will consider neural network operators activated by the ramp function  $R$ , where location and time are very important and effective, with the help of wavelets.

Moreover we will examine and analyse various properties of the wavelet type extension of the neural network operators.

**Definition 2.11. (Wavelet type Neural Network (NN) Operators).** Let  $f : [0, 1] \rightarrow \mathbb{R}$  be a bounded measurable function,  $n \in \mathbb{N}^+$ . and let  $\psi \in L_\infty(\mathbb{R})$  be a father wavelet satisfies **a)-c)**. Then the wavelet type Neural Network (NN) operators activated by the ramp function  $R$ , constructed by using the compactly supported Daubechies wavelets, are defined by:

$$(5) \quad (WN_n f)(t) = n \frac{\sum_{k=0}^n b_R(nt-k) \int_{\mathbb{R}} f(x) \psi(nx-k) dx}{\sum_{k=0}^n b_R(nt-k)}, \quad t \in [0, 1],$$

where  $b_R$  is the bell-shaped function obtained by the ramp function  $R$ .

**Remark 1.** If we choose the father wavelet  $\psi$  as the Haar scaling function, namely  $\psi(x) = \chi_{[0,1]}(x)$ , then clearly our wavelet type operators reduce to the Kantorovich form of the Neural Network (NN) operators. Indeed;

$$\begin{aligned} (WN_n f)(t) &= n \frac{\sum_{k=0}^n b_R(nt-k) \int_{\mathbb{R}} f(x) \psi(nx-k) dx}{\sum_{k=0}^n b_R(nt-k)} \\ &= \frac{\sum_{k=0}^n b_R(nt-k) \int_0^1 f\left(\frac{u+k}{n}\right) \psi(u) dx}{\sum_{k=0}^n b_R(nt-k)}. \end{aligned}$$

This means that our operators constructed by wavelets are a natural extension of the Kantorovich type of the NN operators and also its Durrmeyer type operators.

**Remark 2.** Moreover, the central moments of the wavelet type NN operators (5) are the same as of the classical NN operators (3). Indeed, we get

$$\begin{aligned} (WN_n (x-t)^\beta)(t) &= n \frac{\sum_{k=0}^n b_R(nt-k) \int_{\mathbb{R}} (x-t)^\beta \psi(nx-k) dx}{\sum_{k=0}^n b_R(nt-k)} \\ &= \frac{\sum_{k=0}^n b_R(nt-k) \int_{\mathbb{R}} \left(\frac{u+k}{n} - t\right)^\beta \psi(u) du}{\sum_{k=0}^n b_R(nt-k)} \\ &= \frac{1}{n^\beta} \frac{\sum_{k=0}^n b_R(nt-k) \int_{\mathbb{R}} (u+k-nt)^\beta \psi(u) du}{\sum_{k=0}^n b_R(nt-k)} \\ &= \frac{1}{n^\beta} \frac{\sum_{k=0}^n b_R(nt-k) \int_{\mathbb{R}} \left[ \sum_{i=0}^{\beta} \binom{\beta}{i} u^i (nt-k)^{\beta-i} \right] \psi(u) du}{\sum_{k=0}^n b_R(nt-k)} \end{aligned}$$

Again by the properties of the compactly supported Daubechies wavelets, namely **c)** and **b)**, we get

$$\begin{aligned} (WN_n (x-t)^\beta)(t) &= \frac{1}{n^\beta} \frac{\sum_{k=0}^n b_R(nt-k) (k-nt)^\beta}{\sum_{k=0}^n b_R(nt-k)} \\ &= (N_n (x-t)^\beta)(t). \end{aligned}$$

Throughout this work, for every  $u \in \mathbb{R}$  and for some  $\beta > 0$ , we assume that the algebraic and discrete absolute moment of order  $\beta$  are given by, i.e.,

$$m_\beta(b_R) := \sup_{u \in \mathbb{R}} \sum_{k=0}^n b_R(u-k) (u-k)^\beta,$$

and

$$M_\beta(b_R) := \sup_{u \in \mathbb{R}} \sum_{k=0}^n b_R(u-k) |u-k|^\beta < \infty.$$

### 3. FUNDAMENTAL PROPERTIES AND MAIN RESULTS

We now introduce some notations and structural hypotheses, which will be fundamental in proving our convergence theorems. This section provides the main approximation results of the paper. We have the followings.

We are now ready to establish one of the first main results of this study, which gives a strong relation between NN operators (3) and our new operators (5) constructed by wavelets.:

**Theorem 3.1.** *Let  $f \in B[0,1]$  and let  $\psi \in L_\infty(\mathbb{R})$  be a father wavelet satisfies **a)-c)**. Then the moments of wavelet type NN operators, constructed by using the compactly supported Daubechies wavelets (5) and the NN operators (3) are the same, namely*

$$(WN_n x^s)(t) = (N_n x^s)(t), \quad s = 0, 1, \dots, K$$

holds true.

*Proof.* In view of the definition of the operator (5), we have

$$\begin{aligned} (WN_n x^s)(t) &= \frac{\sum_{k=0}^n \left[ \int_{\mathbb{R}} x^s \psi(nx-k) dx \right] b_R(nt-k)}{\sum_{k=0}^n b_R(nt-k)} \\ &= \frac{\sum_{k=0}^n b_R(nt-k) \int_{\mathbb{R}} \left(\frac{u+k}{n}\right)^s \psi(u) du}{\sum_{k=0}^n b_R(nt-k)} \\ &= \frac{1}{n^s} \frac{\sum_{k=0}^n b_R(nt-k) \int_{\mathbb{R}} (u+k)^s \psi(u) du}{\sum_{k=0}^n b_R(nt-k)} \\ &= \frac{1}{n^s} \frac{\sum_{k=0}^n b_R(nt-k) \int_{\mathbb{R}} \left[ \sum_{i=0}^s \binom{s}{i} u^i k^{s-i} \right] \psi(u) du}{\sum_{k=0}^n b_R(nt-k)}. \end{aligned}$$

In view of **c)**, one has for  $i \neq 0$

$$\int_{\mathbb{R}} \left[ \sum_{i=0}^s \binom{s}{i} u^i k^{s-i} \right] \psi(u) du = 0$$

and for  $i = 0$  and from **b)** we get

$$\begin{aligned} (WN_n x^s)(t) &= \frac{1}{n^s} \frac{\sum_{k=0}^n b_R(nt-k) \int_{\mathbb{R}} k^s \psi(u) du}{\sum_{k=0}^n b_R(nt-k)} \\ &= \frac{\sum_{k=0}^n \frac{k^s}{n^s} b_R(nt-k)}{\sum_{k=0}^n b_R(nt-k)} \\ &= (N_n x^s)(t). \end{aligned}$$

□

**Theorem 3.2.** *Let  $f \in B[0, 1]$  be a measurable function and let  $\psi \in L_\infty(\mathbb{R})$  be a father wavelet satisfies **a)-c)**. Then*

$$\lim_{n \rightarrow \infty} (WN_n f)(t_0) = f(t_0)$$

holds true at each point  $t_0$  of continuity of  $f$ .

*Proof.* Since  $t_0$  is a continuity point of  $f$ , then clearly

$$|f(t) - f(t_0)| < \epsilon$$

holds true when  $|t - t_0| < \delta$ , and

$$|f(t) - f(t_0)| \leq 2 \|f\|_{B[0,1]}$$

holds true, when  $|t - t_0| \geq \delta$ .

So, in view of the definition of the operator (5), one has

$$\begin{aligned} (WS_n f)(t_0) - f(t_0) &= \frac{\sum_{k=0}^n b_R(nt_0 - k) \int_{\mathbb{R}} f(x) \psi(nx - k) dx}{\sum_{k=0}^n b_R(nt_0 - k)} - f(t_0) \\ &= \frac{\sum_{k=0}^n b_R(nt_0 - k) \int_{\mathbb{R}} f\left(\frac{u+k}{n}\right) \psi(u) du}{\sum_{k=0}^n b_R(nt_0 - k)} - f(t_0). \end{aligned}$$

By Theorem 1, we know that

$$(6) \quad (WN_n 1)(t) = (N_n 1)(t) = 1.$$

Hence we can write

$$\begin{aligned} |(WS_n f)(t_0) - f(t_0)| &= \left| \frac{\sum_{k=0}^n b_R(nt_0 - k) \int_{\mathbb{R}} [f\left(\frac{u+k}{n}\right) - f(t_0)] \psi(u) du}{\sum_{k=0}^n b_R(nt_0 - k)} \right| \\ &\leq \frac{\sum_{k=0}^n b_R(nt_0 - k) \int_{\mathbb{R}} |f\left(\frac{u+k}{n}\right) - f(t_0)| |\psi(u)| du}{\sum_{k=0}^n b_R(nt_0 - k)} \end{aligned}$$

Let us divide the last term into two parts as;

$$|(WN_n f)(t_0) - f(t_0)| \leq P_1 + P_2,$$

where

$$P_1 \leq \frac{\sum_{k=0}^n b_R(nt_0 - k) \int_{|\frac{u+k}{n} - t_0| < \delta} |f\left(\frac{u+k}{n}\right) - f(t_0)| |\psi(u)| du}{\sum_{k=0}^n b_R(nt_0 - k)}$$

and

$$P_2 \leq \frac{\sum_{k=0}^n b_R(nt_0 - k) \int_{|\frac{u+k}{n} - t_0| \geq \delta} |f\left(\frac{u+k}{n}\right) - f(t_0)| |\psi(u)| du}{\sum_{k=0}^n b_R(nt_0 - k)}$$

Hence one has

$$\begin{aligned} P_1 &= \frac{\sum_{k=0}^n b_R(nt_0 - k) \int_{|\frac{u+k}{n} - t_0| < \delta} |f\left(\frac{u+k}{n}\right) - f(t_0)| |\psi(u)| du}{\sum_{k=0}^n b_R(nt_0 - k)} \\ &\leq O_R \epsilon \|\psi\|_\infty, \end{aligned}$$

and

$$\begin{aligned}
 P_2 &= \frac{\sum_{k=0}^n b_R(nt_0 - k) \int_{|\frac{u+k}{n}-t_0|\geq\delta} |f(\frac{u+k}{n}) - f(t_0)| |\psi(u)| du}{\sum_{k=0}^n b_R(nt_0 - k)} \\
 &\leq 2 \|f\|_{B[0,1]} \frac{\sum_{k=0}^n b_R(nt_0 - k) \int_{|\frac{u+k}{n}-t_0|\geq\delta} |\psi(u)| du}{\sum_{k=0}^n b_R(nt_0 - k)} \\
 &\leq 2 \|f\|_{B[0,1]} \frac{M_2(b_R)}{\delta^2 n^2} \|\psi\|_\infty = O(n^{-2}).
 \end{aligned}$$

Collecting these estimates we have

$$\lim_{n \rightarrow \infty} (WN_n f)(t_0) = f(t_0).$$

This completes the proof.  $\square$

**Theorem 3.3.** *Let  $f \in C[0, 1]$  and let  $\psi \in L_\infty(\mathbb{R})$  be a father wavelet satisfies **a)-c)**. Then*

$$|(WN_n f)(x) - f(x)| \leq 4\omega(f; 1/n)$$

holds true.

**Corollary 3.4.** *The same arguments of Theorem 2 apply to the case when  $f \in C[0, 1]$ . In this case the convergence is uniform with respect to  $x \in [0, 1]$ , and hence one has*

$$\lim_{n \rightarrow \infty} \|(WN_n f) - f\|_{C[0,1]} = 0.$$

Now, let us consider the following Peetre's  $K$ -functional:

$$(7) \quad K_2(f, \delta) := \inf_{g \in W^2} \{\|f - g\|_{C[0,1]} + \delta \|g''\|_{C[0,1]}\},$$

where  $\delta > 0$  and  $W^2 = \{g \in C[0, 1] : g', g'' \in C[0, 1]\}$ . Then there exists an absolute constant  $C > 0$  such that

$$(8) \quad C^{-1}\omega_2(f, \sqrt{\delta}) \leq K_2(f, \delta) \leq C\omega_2(f, \sqrt{\delta}),$$

where

$$(9) \quad \omega_2(f, \sqrt{\delta}) := \sup_{0 < h \leq \sqrt{\delta}} \sup_{x \in [0,1]} |f(x+2h) - 2f(x+h) + f(x)|$$

is the second order modulus of smoothness of  $f$ . (see [8])

**Theorem 3.5.** *Let  $f \in C[0, 1]$  and let  $\psi \in L_\infty(\mathbb{R})$  be a father wavelet satisfies **a)-c)**. Then*

$$\lim_{n \rightarrow \infty} (WN_n f)(x) = f(x),$$

and

$$|(WN_n f)(x) - f(x)| \leq (K+1) K_2 \left( f; \frac{M_2(b_R) + \lambda^2 O_R + 2\lambda M_1(b_R)}{n^2} \right),$$

where  $K = \lambda \|\psi\|_\infty$  and  $K_2(f; \delta)$  is the Peetre's  $K$ -functional.

*Proof.* Let  $g \in W^2$ . By Taylor's theorem, we have

$$g(t) = g(x) + g'(x)(t-x) + \int_x^t (t-v)g''(v)dv, \quad t \in [0, 1].$$

In view of Remark 2 and (6), applying  $WN_n$  to the both sides of the above equation, we have

$$\begin{aligned}
 |(WN_n g)(x) - g(x)| &= \left| \left( WN_n \left( g'(x)(t-x) + \int_x^t (t-v)g''(v)dv \right) \right) (x) - g(x) \right| \\
 &\leq \frac{\sum_{k=0}^n b_R(nx-k) \int_{\mathbb{R}} \left| \int_x^{\frac{u+k}{n}} \left( \frac{u+k}{n} - v \right) g''(v)dv \right| |\psi(u)| du}{\sum_{k=0}^n b_R(nx-k)} \\
 &\leq \frac{\sum_{k=0}^n b_R(nx-k) \int_0^\lambda \left[ \int_x^{\frac{u+k}{n}} \left| \frac{u+k}{n} - v \right| |g''(v)| dv \right] |\psi(u)| du}{\sum_{k=0}^n b_R(nx-k)} \\
 &\leq \lambda \|\psi\|_\infty \|g''\|_{C[0,1]} \frac{\sum_{k=0}^n b_R(nx-k) \left( \frac{\lambda+k}{n} - x \right)^2}{\sum_{k=0}^n b_R(nx-k)} \\
 &= \lambda \|\psi\|_\infty \|g''\|_{C[0,1]} \frac{\sum_{k=0}^n b_R(nx-k) \left[ \left( \frac{k}{n} - x \right)^2 + \frac{\lambda^2}{n^2} + 2\frac{\lambda}{n} \left( \frac{k}{n} - x \right) \right]}{\sum_{k=0}^n b_R(nx-k)} \\
 &\leq \lambda \|\psi\|_\infty \|g''\|_{C[0,1]} \left[ \frac{M_2(b_R)}{n^2} + \frac{\lambda^2}{n^2} O_R + 2\frac{\lambda}{n} \frac{M_1(b_R)}{n} \right] \\
 &= \frac{\lambda \|\psi\|_\infty \|g''\|_{C[0,1]}}{n^2} [M_2(b_R) + \lambda^2 O_R + 2\lambda M_1(b_R)].
 \end{aligned}$$

Hence, taking infimum on the right hand side over all  $g \in W^2$  and using (7), we get

$$\begin{aligned}
 |(WN_n f)(x) - f(x)| &\leq \inf_{g \in W^2} \{ \|WN_n(f-g)\|_\infty + \|f-g\|_\infty + |(WN_n g)(x) - g(x)| \} \\
 &\leq \inf_{g \in W^2} \left\{ (\lambda \|\psi\|_\infty + 1) \|f-g\|_\infty + \frac{\lambda \|\psi\|_\infty [M_2(b_R) + \lambda^2 O_R + 2\lambda M_1(b_R)]}{n^2} \|g''\|_\infty \right\} \\
 &\leq (K+1) \inf_{g \in W^2} \left\{ \|f-g\|_\infty + \frac{M_2(b_R) + \lambda^2 O_R + 2\lambda M_1(b_R)}{n^2} \|g''\|_\infty \right\} \\
 &= (K+1) K_2 \left( f; \frac{M_2(b_R) + \lambda^2 O_R + 2\lambda M_1(b_R)}{n^2} \right),
 \end{aligned}$$

here  $K = \lambda \|\psi\|_\infty$ . □

**Theorem 3.6.** Let  $f \in C[0, 1]$ ,  $\psi \in L_\infty(\mathbb{R})$  be a father wavelet satisfies **a)-c)** and  $\alpha \in (0, 2)$  be fixed real number. Then

$$\omega_2(f; t) = \mathcal{O}(t^\alpha) \Rightarrow |(WN_n f)(x) - f(x)| = \mathcal{O}(1/n)^\alpha$$

holds true.

*Proof.* In view of (7), (9) and the relation (8) between modulus of smoothness and Peetre's K-functional, we have from Theorem 4

$$\begin{aligned}
 |(WN_n f)(x) - f(x)| &\leq (K+1) K_2 \left( f; \frac{M_2(b_R) + \lambda^2 O_R + 2\lambda M_1(b_R)}{n^2} \right) \\
 &\leq (K+1) C \omega_2 \left( f; \sqrt{\frac{M_2(b_R) + \lambda^2 O_R + 2\lambda M_1(b_R)}{n^2}} \right) \\
 &\leq (K+1) C \left( \frac{M_2(b_R) + \lambda^2 O_R + 2\lambda M_1(b_R)}{n^2} \right)^{\alpha/2}.
 \end{aligned}$$

□



## 4. GRAPHICAL REPRESENTATIONS

Now, we will give some graphical examples for these approach, namely convergence to functions by means of wavelet type Neural Network operators  $(WN_n f)(x)$ .

We note that in all the following Figures, the graph with the red line belongs to the target function.

**Example 4.1.** Let  $f(x) = x^2$ , and take the activation function as a Ramp function for the neural network operators. We consider a special case of the wavelet type Neural Network operators  $(WN_n f)(x)$ , namely Kantorovich type Neural Network operators. Then one has for  $n = 3, 5$  and for  $n = 20$ .

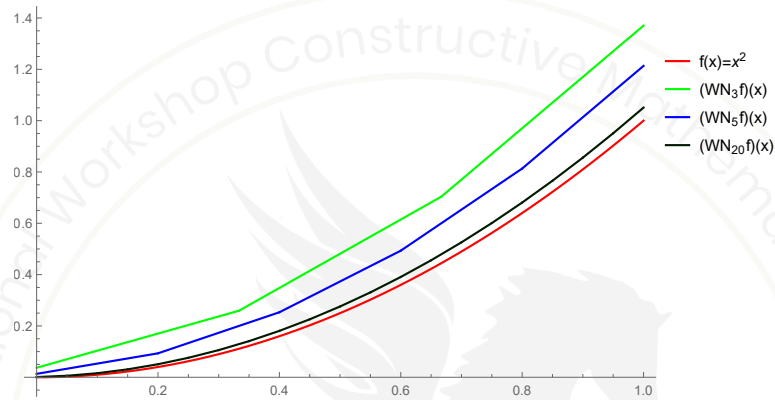


FIGURE 1. Approximation of  $f(x) = x^2$  by Kantorovich type NN operator activated by Ramp function, for  $n = 3, 5$  and  $n = 20$ .

**Example 4.2.** Let  $f(x) = x^2$ , and take the activation function as a Ramp function for the neural network operators. We consider the wavelet type Neural Network operators  $(WN_n f)(x)$  constructed by using Haar scaling function. Then one has for  $n = 3, 5$  and for  $n = 20$ .

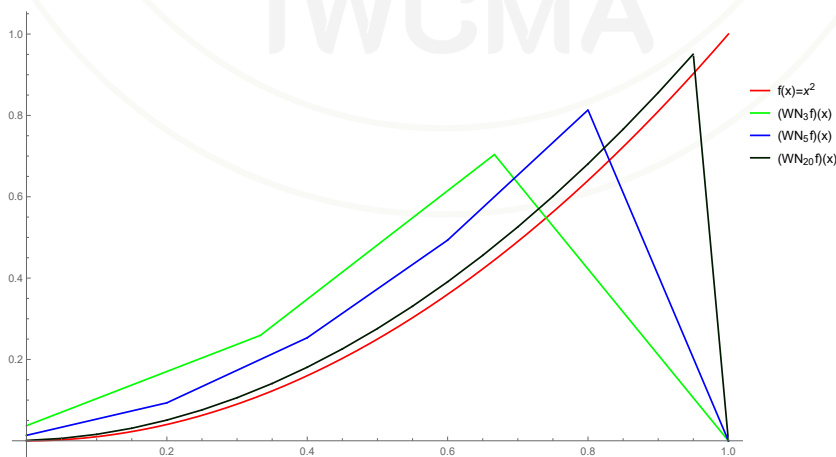


FIGURE 2. Approximation of  $f(x) = x^2$  by Haar Wavelet type NN operator activated by Ramp function, for  $n = 3, 5$  and  $n = 20$ .

**Example 4.3.** Let  $f(x) = x^2$ , and take the activation function as a Ramp function for the neural network operators. We consider the wavelet type Neural Network operators  $(WN_n f)(x)$  constructed by using Shannon wavelet function. Then one has for  $n = 15, 36$  and for  $n = 55$ .

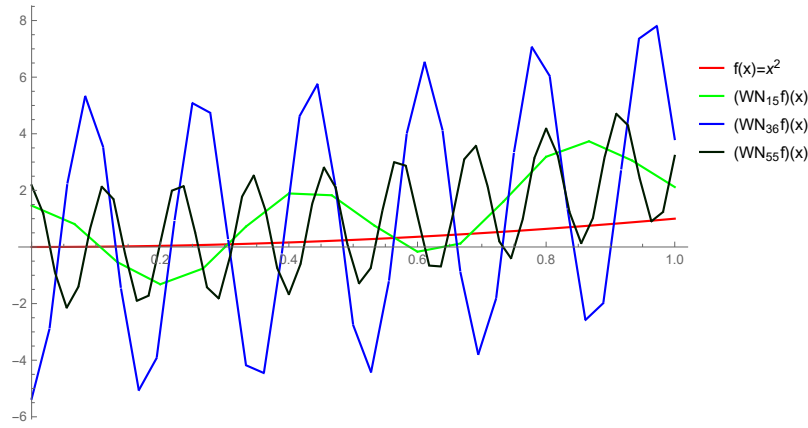


FIGURE 3. Approximation of  $f(x) = x^2$  by Shannon Wavelet type NN operator activated by Ramp function, for  $n = 15, 36$  and  $n = 55$ .

**Example 4.4.** Let  $f(x) = x - x^2$ , and take the activation function as a Ramp function for the neural network operators. We consider a special case of the wavelet type Neural Network operators  $(WN_n f)(x)$ , namely Kantorovich type Neural Network operators. Then one has for  $n = 3, 5$  and for  $n = 20$ .

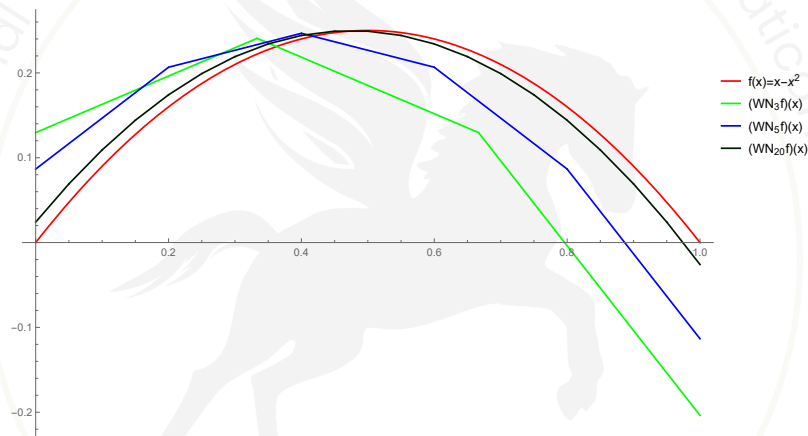


FIGURE 4. Approximation of  $f(x) = x - x^2$  by Kantorovich type NN operator activated by Ramp function, for  $n = 3, 5$  and  $n = 20$ .

**Example 4.5.** Let  $f(x) = x - x^2$ , and take the activation function as a Ramp function for the neural network operators. We consider the wavelet type Neural Network operators  $(WN_n f)(x)$  constructed by using Haar scaling function. Then one has for  $n = 3, 6$  and for  $n = 15$ .

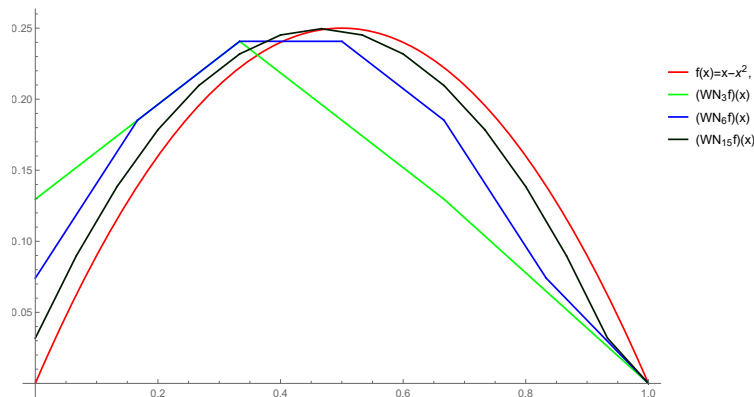


FIGURE 5. Approximation of  $f(x) = x - x^2$  by Haar Wavelet type NN operator activated by Ramp function, for  $n = 3, 6$  and  $n = 15$ .

**Example 4.6.** Let  $f(x) = x - x^2$ , and take the activation function as a Ramp function for the neural network operators. We consider the wavelet type Neural Network operators  $(WN_n f)(x)$  constructed by using Shannon wavelet function. Then one has for  $n = 30, 60$  and for  $n = 150$ .

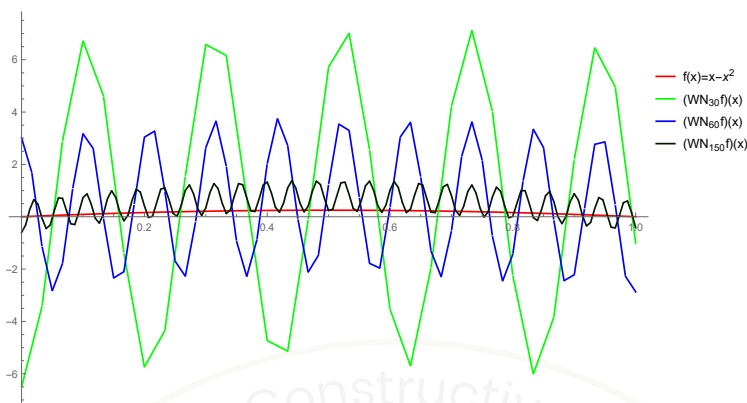


FIGURE 6. Approximation of  $f(x) = x - x^2$  by Shannon Wavelet type NN operator activated by Ramp function, for  $n = 30, 60$  and  $n = 150$ .

**Note 2.** Since the compactly supported Daubechies wavelets are also an unconditional orthonormal base of  $L^p(\mathbb{R})$ , this allows us to investigate the convergence problem on  $L^p(\mathbb{R})$  by means of our wavelet type NN operators (5). These will be the future studies and investigations on this topic.

Moreover, neurocomputing processes deal with multidimensional data, then clearly multivariate Neural Networks have special interest.

Therefore, it is planned as further studies to expand the existing theory to cover the case of multivariate functions and also to make image processing applications.

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# Comparison of imputation and weighting methods in estimation of a finite population mean under random non-response

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**ABSTRACT.** Developing estimators of finite population parameters such as mean, variance and asymptotic mean squared error has been one of the core objectives of sample survey theory and practice. Sample survey practitioners need to assess the properties of these estimators so that better ones can be adopted. In survey sampling, the occurrence of non-response affects inference and optimality of the estimators of finite population parameters. It introduces bias and may cause samples not to follow the the distributions determined by the original sampling design. To compensate for random non-response, imputation methods and weighting techniques can be used. In this paper, a comparison between these two methods of compensating for non-response has been done in two-stage cluster sampling. Simulation results reveal tighter confidence interval lengths, smaller mean squared error values for the estimators developed under the weighting method than its rival estimators obtained using imputation method. Under mild assumptions, the weighting method is shown to be more efficient than the imputation techniques in estimating a finite population mean.

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**KEYWORDS:** Sampling With Replacement, Two-Stage Cluster Sampling, Bias and Mean Squared Error.

## 1. INTRODUCTION

In the estimation of a finite population mean, a lot of significance is attached to efficient and cost-effective survey sampling designs in sample surveys, see for instance, (*cf.* [1]). Careful design of samples based on random selection with known probabilities of population elements should be considered. This gives a target sample of intended respondents where each may provide responses to a set of survey questions that results in an array of responses (*cf.* [23]) observed that non-response occurs if some of the expected responses are missing, for instance where a whole vector of responses is missing for some sampled units or where responses are obtained for some questions and not to others in the sample selected.

The basis for statistical inference is therefore formed by a sampling design that provides a link between a sample and the population. As observed by (*cf.* [10]), a good sample survey practise and efficient methods of compensating for non-response should be adopted.

**1.1. Preliminaries.** In sample surveys, non-response introduces bias in the estimation of a finite population mean. It also causes samples to fail to follow the distributions determined by the original sampling design. The use of regression models is recognized as one of the procedures for reducing bias due to non-response using auxiliary information, for details see (*cf.* [2]). In practise, information on the variables of interest is not available for non-respondents but information on auxiliary variables may be available for non-respondents. To reduce the bias and variance due to non-response, (*cf.* [13]) noted that it is desirable to incorporate auxiliary data into the estimation process where the response probabilities are mostly taken to be correlated with certain characteristics such as age, race and income for a human population survey.

**1.2. Methods of Compensating for Non-response.** Imputation techniques and weighting method are discussed in this paper in the following subsections.

1.2.1. *Imputation Methods.* Imputation entails compensating for non-response values by proxy values as observed by (cf. [20]). Imputation techniques have been used to account for non-response in the study variable in the estimation of finite population mean. For instance, (cf. [21]) applied compromised method of imputation to estimate a finite population mean under two stage cluster sampling, though the method produced a large bias.

Following the procedure by (cf. [9]), let  $\bar{Y} = \frac{1}{N} \sum_1^N Y_i$  represents the finite population mean to be estimated. Besides, let a simple random sample, say,  $S$  with replacement be drawn from the population,  $\theta = 1, 2, \dots, N$  to estimate  $Y$ . The sample  $S$  of units has  $r$  responding units ( $r < n$ ) making a set  $R$  and  $(n - r)$  non-responding units with the subspace  $(n - r)$  having the symbol  $R^C$  in the population. For every unit  $i \in R$ , the value of the survey variable  $y_i$  is obtained.

However, imputed values are to be derived for the non-response set of units, that is, for every  $i \in R^C$ , since the  $y_i$  values are missing. It is assumed that auxiliary data  $x_i$  are known for every  $i \in S$ . The value of the auxiliary,  $x_i$ , imputes the non-response values when  $i \in R^C$ , that is, for a sample  $S$  assume that the data  $x_s = x_i : i \in s$  are given and  $S = R \cup R^C$ . Using this set up, ratio method of imputation, an example of different imputation techniques, can be defined as in equation (2). Using the notations of (cf. [17]), if the  $i^{th}$  unit is to be imputed, the value  $\hat{b}x_i$  is obtained, where  $\hat{b}x_i = \frac{\sum_{i \in R} y_i}{\sum_{i \in R} x_i}$ . The data after imputation becomes

$$(1) \quad y_i = \begin{cases} y_i, & \text{if } i \in R \\ \hat{b}x_i, & \text{if } i \in R^C \end{cases}$$

For details, see (cf. [21]). The imputation ratio estimator is given by

$$(2) \quad \bar{y}_{RAT} = \bar{y}_r \frac{\bar{x}_n}{\bar{x}_r}$$

where  $\bar{x}_n = \frac{1}{n} \sum_{i \in s} x_i$ ,  $\bar{x}_r = \frac{1}{r} \sum_{i \in R} x_i$  and  $\bar{y}_r = \frac{1}{r} \sum_{i \in R} y_i$ . The bias and the mean squared error of the imputation ratio estimator due to (cf. [21]) are given below.

$$(3) \quad B(\bar{y}_{RAT}) = \left( \frac{1}{r} - \frac{1}{n} \right) \bar{Y} (C_x^2 - \rho C_x C_y)$$

where  $C_x = \frac{s_x}{\bar{X}}$ ,  $C_y = \frac{s_y}{\bar{Y}}$  and  $\rho = \frac{S_{xy}}{S_x S_y}$ .  $S_x$  and  $S_y$  are the standard deviations of  $X$  and  $Y$  values respectively while  $S_{xy}$  is the co-variance between  $X$  and  $Y$ . The MSE is given by

$$(4) \quad MSE(\bar{y}_{RAT}) = \left( \frac{1}{n} - \frac{1}{N} \right) S_y^2 + \left( \frac{1}{r} - \frac{1}{n} \right) \left[ S_y^2 + R_1^2 S_x^2 - 2R_1 S_{xy} \right]$$

where  $R_1 = \frac{\bar{Y}}{\bar{X}}$ . (cf. [1]) observed that though this method of imputation is better than other existing techniques like mean and compromised method of imputation, its bias and MSE are still large compared to rival approaches of compensating for non-response such as weighting techniques.

1.2.2. *Weighting Method.* It has been observed by authors such as (cf. [16]) that non-response causes loss of observations and therefore weighting means that the weights are increased for all or almost all of the elements that fail to respond in a survey. For instance, (cf. [4]) and (cf. [11]) discussed a modified Horvitz-Thompson estimator to correct for non-response problem using the weighting strategy. The estimator used was defined by

$$(5) \quad \hat{y}_{HT} = N^{-1} \sum_{k=1}^N (\phi_k \hat{p}_{HT})^{-1} y_k \tau_k$$

where  $\hat{p}_{HT}$  is given by

$$\hat{p}_{HT} = N^{-1} \sum_{k=1}^N \phi_k^{-1} \gamma_k$$

where  $y_k, k \in U$  is the value of the  $k^{th}$  survey variable taken from a sample  $s$  selected from a finite population,  $U = (1, 2, \dots, N)$ ,  $\phi_k$  is the inclusion probability given by  $\phi_k = p_r(k \in s)$ ,  $\gamma_k$  is the value of the  $k^{th}$  respondent in the sample selected,  $s$ . The estimator  $\hat{y}_{HT}$  adjusts the



weights by an unbiased estimator  $\hat{p}_{HT}$ , of the response probabilities of the population mean which is given by

$$(6) \quad \bar{P}_N = N^{-1} \sum_{i=1}^N p_i$$

Thus an approximate bias of the estimator  $\hat{y}_{HT}$  is given by

$$(7) \quad Bias(\hat{y}_{HT}) = E(\hat{y}_{HT}) - \bar{Y} = \bar{Y}^* - \bar{Y} = \hat{P}_N^{-1} C_p Y$$

where

$$(8) \quad \bar{Y}^* = N^{-1} P_N^{-1} \sum_{i=1}^N p_i y_i$$

and

$$(9) \quad C_p Y = N^{-1} \sum_{i=1}^N (p_i - \bar{P}_N)(y_i - \bar{Y}).$$

If  $C_p Y$  is close to zero, the bias will be small, for more details see (cf. [20]). The adjusted Horvitz-Thompson estimator,  $\hat{y}_{HT}$ , is an illustration of re-weighting measurements of respondents without using auxiliary information. However, in this paper, auxiliary information is used in the estimation procedure.

The population mean,  $\bar{Y} = \frac{1}{N} \sum_1^N Y_i$ , is estimated by selecting a sample of size  $n$  at random with replacement. If the responding units of item  $y$  are independent so that the probability of unit  $j$  responding in cluster  $i$  is  $p_{ij}(i = 1, 2, \dots, n; j = 1, 2, \dots, m)$ , then following the work of (cf. [18]), a weighted estimator,  $\hat{y}_I$ , for  $\bar{Y}$  is given by

$$(10) \quad \hat{y}_I = \frac{1}{\sum_{i,j \in s} w_{ij}} \left[ \sum_{i,j \in s_r} w_{ij} y_{ij} + \sum_{i,j \in s_m} w_{ij} y_{ij}^* \right]$$

where  $w_{ij} = \frac{1}{\pi_{ij}}$  gives the survey weight tied to unit  $j$  in cluster  $i$  and  $\pi_{ij} = p[i, j \in s]$  is its probability of inclusion,  $s_r$ , is the set of  $r$  responding units to item  $y$ ,  $s_m$  is the set of  $m$  units that failed to respond to item  $y$  so that  $r + m = n$  while  $y_{ij}^*$  is the value imputed so that the missing value  $y_{ij}$  is compensated for, (cf. [2]).

## 2. MAIN RESULTS

**2.1. The proposed Estimator of Finite Population Mean.** Consider a finite population of size  $N$  consisting of  $M$  clusters with  $N_j$  elements in the  $j^{th}$  cluster. A sample of  $m$  clusters is selected so that  $n_{1i}$  units respond and  $n_{2i}$  units fail to respond. Let  $y_{ij}$  denote the value of the survey variable  $y$  for unit  $j$  in cluster  $i$ , for  $i = 1, 2, \dots, N, j = 1, 2, \dots, N_i$  and let population mean be given by

$$(11) \quad \bar{\bar{Y}} = \frac{1}{N} \sum_{i=1}^M \sum_{j=1}^{N_i} Y_{ij}.$$

The proposed estimator is given by

$$(12) \quad \hat{\bar{Y}}_{INW} = \frac{1}{M} \left\{ \frac{1}{n_{1i}} \sum_{i \in s} \sum_{j \in s} y_{ij} + \frac{1}{n_{2i}} \sum_{i \in s} \sum_{j \notin s} \hat{y}_{ij} \right\}.$$

where  $\hat{y}_{ij}$  is an estimator of the non-response component of the sample. Assuming auxiliary information,  $X_{ij}$ , is known throughout,  $\hat{y}_{ij}$  can be obtained using the improved Nadaraya-Watson regression technique by

$$(13) \quad \hat{y}_{ij} = m_{INW}(\hat{x}_{ij}) = \frac{\sum_{i \in s} \sum_{j \in s} \frac{1}{\lambda_{ij}} K\left(\frac{x - X_{ij}}{\lambda_{ij} b}\right) Y_{ij}}{\sum_{i \in s} \sum_{j \in s} \frac{1}{\lambda_{ij}} K\left(\frac{x - X_{ij}}{\lambda_{ij} b}\right)}$$

so that the estimator of finite population mean can be re-written as

$$(14) \quad \hat{\bar{Y}}_{INW} = \frac{1}{M} \left\{ \frac{1}{n_{1i}} \sum_{i \in s} \sum_{j \in s} y_{ij} + \frac{1}{n_{2i}} \sum_{i \in s} \sum_{j \notin s} m_{INW}(\hat{x}_{ij}) \right\}.$$

A special case where  $n_{1i} = n_{2i} = n$  is assumed in this study. This simplifies mathematical computations so that equation (15) can be re-written as

$$(15) \quad \hat{\bar{Y}}_{INW} = \frac{1}{Mn} \left\{ \sum_{i \in s} \sum_{j \in s} y_{ij} + \sum_{i \in s} \sum_{j \notin s} m_{INW}(\hat{x}_{ij}) \right\},$$

where  $m_{INW}(\hat{x}_{ij})$  is the improved Nadaraya-Watson kernel regression estimator given in (??), which is a weighted sum of the values of the survey variable  $Y_{ij}$ 's. Data is generated using a regression model given by

$$(16) \quad \hat{Y}_{ij} = m(\hat{x}_{ij}) + \hat{e}_{ij}$$

where  $m(\cdot)$  is an unknown smooth function of auxiliary random variables,  $X_{ij}$ . It is assumed that the error term,  $\hat{e}_{ij}$ , satisfies the following conditions:

$$(17) \quad E(\hat{e}_{ij}) = 0, Var(\hat{e}_{ij}) = \sigma_{ij}^2, Cov(\hat{e}_i, \hat{e}_j) = 0, \text{ for } i \neq j$$

Hence the unspecified function of the auxiliary random variables,  $m(\hat{x}_{ij})$ , is replaced by the improved Nadaraya-Watson kernel estimator,  $m_{INW}(\hat{x}_{ij})$ . The estimator can be re-written as

$$(18) \quad m_{INW}(\hat{x}_{ij}) = \sum_{i \in s} \sum_{j \in s} w(x_{ij}) Y_{ij}.$$

where  $w(x_{ij}) = \frac{\frac{1}{\lambda_{ij}} K\left(\frac{x-X_{ij}}{\lambda_{ij}^b}\right)}{\sum_{i \in s} \sum_{j \in s} \frac{1}{\lambda_{ij}} K\left(\frac{x-X_{ij}}{\lambda_{ij}^b}\right)}$  are the improved Nadaraya-Watson kernel weights

where  $K(\cdot)$  is a given kernel function assumed to be symmetrical;  $b$  is a smoothing parameter while  $\lambda_{ij}$  is the local bandwidth given by

$$(19) \quad \lambda_{ij} = \{m(X_{ij}/a)\}^{-\alpha}.$$

where  $\alpha$  is a sensitivity parameter which satisfies  $0 \leq \alpha \leq 1$ . It has been suggested by (cf. [12]) that taking  $\alpha = \frac{1}{2}$  produce good results. Since the choice of the kernel function is not critical for the performance of the kernel regression estimator, a simplified Gaussian kernel with mean 0 and variance 1 is used in this study. This is given by

$$(20) \quad K(w) = \frac{1}{\sqrt{2\pi}} e^{-\left(\frac{w^2}{2}\right)} = \frac{1}{\sqrt{2\pi}} e^{-\left(\frac{\left(\frac{x-X_{ij}}{\lambda_{ij}^b}\right)^2}{2}\right)}$$

In this case, the improved Nadaraya-Watson kernel estimation at any point  $x_{ij}$  is given by

$$(21) \quad \hat{y}_{ij} = m_{INW}(\hat{x}_{ij}) = \frac{\sum_i \sum_j \frac{1}{\sqrt{2\pi}} e^{-\left(\frac{\left(\frac{x-X_{ij}}{\lambda_{ij}^b}\right)^2}{2}\right)} Y_{ij}}{\sum_i \sum_j \frac{1}{\sqrt{2\pi}} e^{-\left(\frac{\left(\frac{x-X_{ij}}{\lambda_{ij}^b}\right)^2}{2}\right)}}$$

where  $b$  is the bandwidth while  $\lambda_{ij}$  is given in equation (??) due to (cf. [7]).

This provides a way of estimating the non-response values of the survey variable  $Y_{ij}$ , in the  $i^{th}$  cluster given the auxiliary values  $x_{ij}$ , for a specified kernel function.

**2.2. The Asymptotic Bias of the Proposed Estimator.** The expected value of the proposed estimator is given by

$$(22) \quad E(\hat{\bar{Y}}_{INW}) = \frac{1}{Mn} \left\{ \sum_{i=1}^n \sum_{j=1}^m Y_{ij} + \sum_{i=n+1}^N \sum_{j=m+1}^M m_{INW}(\hat{x}_{ij}) \right\}$$

Re-writing equation (13) using the property of symmetry associated with Nadaraya-Watson estimator,

$$(23) \quad m_{INW}(\hat{x}_{ij}) = \frac{\sum_{i \in s} \sum_{j \in s} K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij} b}\right) Y_{ij}}{\sum_{i \in s} \sum_{j \in s} K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij} b}\right)}, i = 1, 2, \dots, n; j = 1, 2, \dots, m$$

Following the procedure by (cf. [22]), equation (23) can be re-written as

$$(24) \quad m_{INW}(\hat{x}_{ij}) = \frac{1}{g(\hat{x}_{ij})} \left[ \frac{1}{mn(\lambda_{ij} b)} \sum_i \sum_j K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij} b}\right) Y_{ij} \right]$$

where  $g(\hat{x}_{ij})$  is the estimated marginal density of auxiliary variables  $X_{ij}$ . The bias of the estimator can be written as

$$(25) \quad Bias(\hat{\bar{Y}}_{INW}) = E(\hat{\bar{Y}}_{INW} - \bar{Y})$$

$$(26) \quad Bias(\hat{\bar{Y}}_{INW}) = E \left\{ \frac{1}{Mn} \left[ \sum_{i=1}^n \sum_{j=1}^m Y_{ij} + \sum_{i=n+1}^N \sum_{j=m+1}^M m_{INW}(\hat{x}_{ij}) \right] - \frac{1}{Mn} \left[ \sum_{i=1}^n \sum_{j=1}^m Y_{ij} + \sum_{i=n+1}^N \sum_{j=m+1}^M Y_{ij} \right] \right\}$$

which reduces to

$$(27) \quad Bias(\hat{\bar{Y}}_{INW}) = \frac{1}{Mn} E \left\{ \sum_{i=n+1}^N \sum_{j=m+1}^M m_{INW}(\hat{x}_{ij}) - \sum_{i=n+1}^N \sum_{j=m+1}^M Y_{ij} \right\}$$

Re-writing the regression model given by  $Y_{ij} = m(X_{ij}) + e_{ij}$  as

$$(28) \quad Y_{ij} = m(x_{ij}) + [m(X_{ij}) - m(x_{ij})] + e_{ij}$$

and substituting it in equation (24) gives

$$(29) \quad m_{INW}(\hat{x}_{ij}) = \frac{1}{g(\hat{x}_{ij})} \left[ \frac{1}{mn(\lambda_{ij} b)} \sum_i \sum_j K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij} b}\right) (m(x_{ij}) + [m(X_{ij}) - m(x_{ij})] + e_{ij}) \right]$$

Hence the first term in equation (27) before taking expectation is given as:

$$(30) \quad \begin{aligned} & \frac{1}{Mn} \left\{ \frac{\frac{1}{mn} \sum_{i=n+1}^N \sum_{j=m+1}^M K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij} b}\right) Y_{ij}}{g(\hat{x}_{ij})} \right\} \\ &= \frac{1}{Mng(\hat{x}_{ij})} \left\{ \frac{1}{mn(\lambda_{ij} b)} \sum_{i=n+1}^N \sum_{j=m+1}^M K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij} b}\right) m(x_{ij}) \right. \\ &+ \frac{1}{mn(\lambda_{ij} b)} \sum_{i=n+1}^N \sum_{j=m+1}^M K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij} b}\right) [m(X_{ij}) - m(x_{ij})] \\ &+ \left. \frac{1}{mn(\lambda_{ij} b)} \sum_{i=n+1}^N \sum_{j=m+1}^M K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij} b}\right) e_{ij} \right\} \end{aligned}$$

Simplifying equation (30), the following is obtained:

$$(31) \quad \begin{aligned} & \frac{1}{Mn} \left\{ \frac{1}{mn(\lambda_{ij}b) g(\hat{x}_{ij})} \sum_{i=n+1}^N \sum_{j=m+1}^M K \left( \frac{X_{ij} - x_{ij}}{\lambda_{ij}b} \right) Y_{ij} \right\} \\ &= \frac{1}{Mn} \left( \frac{1}{mn(\lambda_{ij}b) g(\hat{x}_{ij})} \right) \left\{ \sum_{i=n+1}^N \sum_{j=m+1}^M g(\hat{x}_{ij}) m(x_{ij}) \right. \\ & \quad \left. + m_1(\hat{x}_{ij}) + m_2(\hat{x}_{ij}) \right\}, \end{aligned}$$

where

$$(32) \quad m_1(\hat{x}_{ij}) = \frac{1}{mn(\lambda_{ij}b)} \sum_{i=n+1}^N \sum_{j=m+1}^M K \left( \frac{X_{ij} - x_{ij}}{\lambda_{ij}b} \right) [m(X_{ij}) - m(x_{ij})].$$

$$(33) \quad m_2(\hat{x}_{ij}) = \frac{1}{mn(\lambda_{ij}b)} \sum_{i=n+1}^N \sum_{j=m+1}^M K \left( \frac{X_{ij} - x_{ij}}{\lambda_{ij}b} \right) e_{ij}.$$

Taking conditional expectation of equation (31) leads to

$$(34) \quad \begin{aligned} E \left[ \sum_{i=n+1}^N \sum_{j=m+1}^M m_{INW}(\hat{x}_{ij})/x_{ij} \right] &= \frac{1}{Mn} E \left[ \frac{1}{mn(\lambda_{ij}b)} \sum_{i=n+1}^N \sum_{j=m+1}^M \left[ m(x_{ij}) \right. \right. \\ & \quad \left. \left. + \frac{m_1(\hat{x}_{ij})}{g(\hat{x}_{ij})} + \frac{m_2(\hat{x}_{ij})}{g(\hat{x}_{ij})} \right] \right]. \end{aligned}$$

The following theorem due to (*cf.* [8]) and applied by (*cf.* [19]) was used in obtaining asymptotic bias and variance of the estimator using conditional expectations.

**Theorem 2.1.** *Let  $K(w)$  be a symmetric density function with  $\int wk(w)dw = 0$  and  $\int w^2k(w)dw = k_2$ . Assume  $n$  and  $N$  increase together such that  $\frac{n}{N} \rightarrow \pi$  with  $0 < \pi < 1$ . Besides, assume the sampled and non-sampled values of  $x$  are in the interval  $[c, d]$  and are obtained by densities  $d_s$  and  $d_{p-s}$  respectively where both are bounded away from zero on  $[c, d]$  with continuous second derivatives. If for any variable  $Z$ ,  $E(Z/U = u) = A(u) + O(B)$  and  $Var(Z/U = u) = O(C)$ , then  $Z = A(u) + O_p(B + C^{\frac{1}{2}})$ .*

Using this theorem, the asymptotic bias can further be derived and simplified. From the conditions of the error term stated in (17), it follows that  $E(e_{ij}/X_{ij}) = 0$ . Therefore,  $E[m_2(\hat{x}_{ij})] = 0$ . Thus,  $E[m_1(\hat{x}_{ij})]$  can be obtained as follows:

$$(35) \quad \begin{aligned} E \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] &= \frac{1}{Mn} \left( \frac{1}{mn(\lambda_{ij}b)} \right) E \left\{ \sum_{i=n+1}^N \sum_{j=m+1}^M K \left( \frac{X_{ij} - x_{ij}}{\lambda_{ij}b} \right) \right. \\ & \quad \left. \times [m(X_{ij}) - m(x_{ij})] \right\}. \end{aligned}$$

Using substitution and change of variable technique given by

$$(36) \quad \left. \begin{aligned} w &= \frac{V - x_{ij}}{\lambda_{ij}b} \\ V &= x_{ij} + (\lambda_{ij}b)w \\ dV &= (\lambda_{ij}b)dw \end{aligned} \right\}$$

Equation (35) can be simplified to:

$$(37) \quad \begin{aligned} E \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] &= \frac{1}{Mn} \left\{ \frac{Mn - mn}{mn} \int k(w) [m(x_{ij} + (\lambda_{ij}b)w) \right. \\ & \quad \left. - m(x_{ij})] \int g(x_{ij} + (\lambda_{ij}b)w) dw \right\} \end{aligned}$$

Using Taylor's series expansion about the point  $x_{ij}$ , the  $k^{th}$  order kernel can be derived as follows:

$$(38) \quad \begin{aligned} g(x_{ij} + (\lambda_{ij}b)w) &= g(x_{ij}) + g'(x_{ij})(\lambda_{ij}b)w + \frac{1}{2}g''(x_{ij})(\lambda_{ij}b)^2w^2 + \dots \\ &+ \frac{1}{k!}g^k(x_{ij})(\lambda_{ij}b)^kw^k + o((\lambda_{ij}b)^2). \end{aligned}$$

Similarly,

$$(39) \quad \begin{aligned} m(x_{ij} + (\lambda_{ij}b)w) &= m(x_{ij}) + m'(x_{ij})(\lambda_{ij}b)w + \frac{1}{2}m''(x_{ij})(\lambda_{ij}b)^2w^2 + \dots \\ &+ \frac{1}{k!}m^k(x_{ij})(\lambda_{ij}b)^kw^k + o((\lambda_{ij}b)^2). \end{aligned}$$

Therefore, expanding equation (37) up to order  $o((\lambda_{ij}b)^2)$  and simplifying gives

$$(40) \quad \begin{aligned} E \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] &= \frac{1}{Mn} \left\{ \left( \frac{Mn - mn}{mn} \right) g(x_{ij})m'(x_{ij})(\lambda_{ij}b) \int w k(w)dw \right. \\ &+ \left( \frac{Mn - mn}{mn} \right) g'(x_{ij})m'(x_{ij})(\lambda_{ij}b)^2 \int w^2 k(w)dw \\ &+ \left( \frac{Mn - mn}{mn} \right) \frac{1}{2} g(x_{ij})m''(x_{ij})(\lambda_{ij}b)^2 \\ &\left. \times \int w^2 k(w)dw + o((\lambda_{ij}b)^2) \right\}. \end{aligned}$$

Using the conditions due to (cf. [8]) given by  $\int_{-\infty}^{\infty} k(w)dw = 1$ ,  $\int_{-\infty}^{\infty} wk(w)dw = 0$  and  $\int_{-\infty}^{\infty} w^2k(w)dw = d_k$ , the derivation in equation (40) can further be simplified to obtain:

$$(41) \quad \begin{aligned} E \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] &= \frac{1}{Mn} \left( \frac{Mn - mn}{mn} \right) \left[ g'(x_{ij})m'(x_{ij}) \right. \\ &\left. + \frac{1}{2}g(x_{ij})m''(x_{ij}) \right] (\lambda_{ij}b)^2 d_k + o((\lambda_{ij}b)^2). \end{aligned}$$

Hence the expected value of the second term in equation (34) then becomes:

$$(42) \quad \begin{aligned} E \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] &= \frac{1}{Mn} \left\{ \left( \frac{Mn - mn}{mn} \right) \left[ \frac{1}{2g(\hat{x}_{ij})} m''(x_{ij})g(x_{ij}) \right. \right. \\ &\left. \left. + \frac{g'(x_{ij})m'(x_{ij})}{g(\hat{x}_{ij})} \right] (\lambda_{ij}b)^2 d_k + o((\lambda_{ij}b)^2) \right\}. \end{aligned}$$

Simplifying equation (42) gives:

$$(43) \quad E \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] = \frac{1}{Mn} \left\{ \left( \frac{Mn - mn}{mn} \right) (\lambda_{ij}b)^2 d_k C(x) + o((\lambda_{ij}b)^2) \right\},$$

where  $C(x) = [g(\hat{x}_{ij})]^{-1} \left[ \frac{1}{2}m''(x_{ij})g(x_{ij}) + g'(x_{ij})m'(x_{ij}) \right]$  and  $d_k = \int w^2k(w)dw$ .

Using equation of the bias given in (25) and the conditional expectation in equation (34), the following equation for the conditional bias of the estimator was obtained:

$$(44) \quad Bias(\hat{\bar{Y}}_{INW}/x_{ij}) = \frac{1}{Mn} \left\{ \left( \frac{Mn - mn}{mn} \right) (\lambda_{ij}b)^2 d_k C(x) + o((\lambda_{ij}b)^2) \right\}.$$

In the next subsection, the asymptotic variance of the estimator is also derived.

**2.3. Asymptotic Variance of the Proposed Estimator.** Using equation (15), the conditional variance of the estimator is given as

$$(45) \quad Var(\hat{\bar{Y}}_{INW}/x_{ij}) = Var \left\{ \frac{1}{Mn} \left\{ \sum_{i=1}^n \sum_{j=1}^m Y_{ij} + \sum_{i=n+1}^N \sum_{j=m+1}^M m_{INW}(\hat{x}_{ij}) \right\} \right\}$$

$$(46) \quad = \left( \frac{1}{Mn} \right)^2 Var \left\{ \sum_{i=1}^n \sum_{j=1}^m Y_{ij} + \sum_{i=n+1}^N \sum_{j=m+1}^M m_{INW}(\hat{x}_{ij}) \right\},$$

where  $m_{INW}(\hat{x}_{ij})$  is given by

$$(47) \quad m_{INW}(\hat{x}_{ij}) = \frac{1}{g(\hat{x}_{ij})} \left[ \frac{1}{mn(\lambda_{ij}b)} \sum_i \sum_j K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij}b}\right) Y_{ij} \right],$$

where  $g(\hat{x}_{ij}) = \frac{1}{mn(\lambda_{ij}b)} \sum_i \sum_j K\left(\frac{x - X_{ij}}{\lambda_{ij}b}\right)$  is the estimated marginal density of auxiliary variables  $X_{ij}$ , for details see (cf. [15]). Re-writing the regression model  $Y_{ij} = m(X_{ij}) + e_{ij}$  as  $Y_{ij} = m(x_{ij}) + [m(X_{ij}) - m(x_{ij})] + e_{ij}$  and substituting in equation (47) leads to

$$(48) \quad \begin{aligned} Var(\hat{\bar{Y}}_{INW}/x_{ij}) &= \left(\frac{1}{Mn}\right)^2 Var \left\{ \sum_{i=1}^n \sum_{j=1}^m Y_{ij} \right. \\ &\quad \left. + \left(\frac{1}{mn(\lambda_{ij}b)g(\hat{x}_{ij})}\right) \left\{ \sum_{i=n+1}^N \sum_{j=m+1}^M g(\hat{x}_{ij})m(x_{ij}) + m_1(\hat{x}_{ij}) + m_2(\hat{x}_{ij}) \right\} \right\}. \end{aligned}$$

From equation (33),

$$(49) \quad m_2(\hat{x}_{ij}) = \frac{1}{mnb} \sum_{i=n+1}^N \sum_{j=m+1}^M K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij}b}\right) e_{ij}$$

Hence

$$(50) \quad Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_2(\hat{x}_{ij})] = \frac{1}{(Mn)^2} \left(\frac{Mn - mn}{mn(\lambda_{ij}b)}\right)^2 \sum_{i=1}^n \sum_{j=1}^m Var(D_x),$$

where  $D_x = K\left(\frac{X_{ij} - x_{ij}}{\lambda_{ij}b}\right) e_{ij}$ . Expressing equation (50) in terms of expectation the following equation is obtained

$$(51) \quad Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_2(\hat{x}_{ij})] = \frac{1}{(Mn)^2} \left[\frac{(Mn - mn)^2}{mn(\lambda_{ij}b)^2}\right] \left\{ E[D_x]^2 - [E(D_x)]^2 \right\}.$$

Using the fact that the conditional expectation  $E(e_{ij}/X_{ij}) = 0$ , the second term in equation (51) reduces to zero. Therefore,

$$(52) \quad Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_2(\hat{x}_{ij})] = \frac{1}{(Mn)^2} \left[\frac{(Mn - mn)^2}{mn(\lambda_{ij}b)^2}\right] \sigma_{ij}^2,$$

where  $E(e_{ij}/X_{ij})^2 = \sigma_{ij}^2$ .

Let  $X = X_{ij}$ , and  $x = x_{ij}$  and make the following substitutions

$$(53) \quad \left. \begin{aligned} w &= \frac{X-x}{\lambda_{ij}b} \\ X - x &= (\lambda_{ij}b)w \\ dX &= (\lambda_{ij}b)dw. \end{aligned} \right\}$$

so that

$$(54) \quad Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_2(\hat{x}_{ij})] = \frac{(Mn - mn)^2}{mn(\lambda_{ij}b)^2(Mn)^2} \int K\left(\frac{X - x}{\lambda_{ij}b}\right)^2 \sigma_x^2 g(X) dX$$

Using the change of variables technique and simplifying, equation (54) reduces to

$$(55) \quad Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_2(\hat{x}_{ij})] = \frac{(Mn - mn)^2}{mn(\lambda_{ij}b)(Mn)^2} \int K(w)^2 \sigma_x^2 g(x) dw + o\left(\frac{1}{mn(\lambda_{ij}b)}\right).$$



Following the same procedure for getting the variance of  $m_2(\hat{x}_{ij})$ ,

$Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})]$  can similarly be obtained as follows:

$$(56) \quad Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] = \frac{1}{(Mn)^2} Var \sum_{i=n+1}^N \sum_{j=m+1}^M \left[ \frac{1}{mn(\lambda_{ij}b)} \sum_{i=1}^n \sum_{j=1}^m K \left( \frac{X_{ij} - x_{ij}}{\lambda_{ij}b} \right) \right] \\ \times [m(X_{ij}) - m(x_{ij})]$$

Equation (56) can be re-written as

$$(57) \quad Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] = \frac{(Mn - mn)^2}{mnb^2(Mn)^2} Var K \left( \frac{X_{ij} - x_{ij}}{\lambda_{ij}b} \right)^2 [m(X_{ij}) \\ - m(x_{ij})]^2 g(X) dX$$

where  $X = (\lambda_{ij}b)w + x$  so that  $dX = (\lambda_{ij}b)dw$ . Changing variables and applying Taylor's series expansion about the point  $x_{ij}$  leads to

$$(58) \quad Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] = \frac{(Mn - mn)^2}{mn(\lambda_{ij}b)^2(Mn)^2} \int K(w^2) [m(x + (\lambda_{ij}b)w) - m(x)]^2 \\ \times g(x + (\lambda_{ij}b)w) dw$$

which gives

$$(59) \quad Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] = \frac{(Mn - mn)^2}{mn(\lambda_{ij}b)^2(Mn)^2} \int K(w^2) [m(x) + m'(x)(\lambda_{ij}b)w \\ + \dots - m(x)]^2 (g(x) + g'(x)(\lambda_{ij}b)w) dw.$$

Following the procedure by (cf. [3]) and simplifying, equation (59) reduces to

$$(60) \quad Var \sum_{i=n+1}^N \sum_{j=m+1}^M [m_1(\hat{x}_{ij})] = o \left[ \frac{(Mn - mn)^2 b^2}{mn(\lambda_{ij}b)} \right].$$

For large samples, as  $n \rightarrow N, m \rightarrow M$  and  $b \rightarrow 0$ , then  $mn(\lambda_{ij}b) \rightarrow \infty$ . Hence the variance in equation (59) asymptotically tends to zero, i.e,  $Var \sum_{i=1}^N \sum_{j=1}^M [m_1(\hat{x}_{ij})] \rightarrow 0$  so that the variance of the estimator of the population mean reduces to

$$(61) \quad Var \left( \hat{\bar{Y}}_{INW} / x_{ij} \right) = \frac{(Mn - mn)^2}{mn(\lambda_{ij}b)(Mn)^2} \sum_{i=n+1}^N \sum_{j=m+1}^M Var \left[ m(x_{ij}) + \frac{m_1(\hat{x}_{ij}) + m_2(\hat{x}_{ij})}{g(\hat{x}_{ij})} \right].$$

Simplifying equation (61) leads to

$$(62) \quad Var \left( \hat{\bar{Y}}_{INW} / x_{ij} \right) = \frac{(Mn - mn)^2}{mn(\lambda_{ij}b)(Mn)^2 [g(\hat{x}_{ij})]^2} Var \left\{ \sum_{i=n+1}^N \sum_{j=m+1}^M [m_2(\hat{x}_{ij})] \right\}.$$

Substituting equation (55) into (62) yields the following:

$$(63) \quad Var \left( \hat{\bar{Y}}_{INW} / x_{ij} \right) = \frac{1}{(Mn)^2} \left\{ \frac{(Mn - mn)^2 \int K(w)^2 \sigma_{x_{ij}}^2 dw}{mn(\lambda_{ij}b)g(\hat{x}_{ij})} + o \left[ \frac{(Mn - mn)^2}{mn(\lambda_{ij}b)} \right. \right. \\ \left. \left. + \frac{1}{mn(\lambda_{ij}b)} \right] \right\}.$$

**2.4. Mean Squared Error of the Proposed Estimator.** The conditional *MSE* of the estimator of finite population mean combines the conditional squared bias and the conditional variance of the estimator, that is,

$$(64) \quad MSE(\hat{\bar{Y}}_{INW} / x_{ij}) = Var(\hat{\bar{Y}}_{INW} / x_{ij}) + Bias^2(\hat{\bar{Y}}_{INW} / x_{ij})$$

Which on simplification leads to

$$(65) \quad \begin{aligned} MSE(\hat{\bar{Y}}/X_{ij} = x_{ij}) &= \frac{1}{(Mn)^2} \left\{ \frac{(Mn - mn)^2 H(w) \sigma_{x_{ij}}^2}{mn(\lambda_{ij}b)g(\hat{x}_{ij})} \right. \\ &+ \left[ \frac{(Mn - mn)^2}{4(mn)^2(Mn^2)} (\lambda_{ij}b)^2 d_k^2 \left[ m\mu(x_{ij})g(x_{ij}) + \frac{2g'(x_{ij})m'(x_{ij})}{g(\hat{x}_{ij})} \right]^2 \right. \\ &\left. \left. + o\left( \frac{1}{Mn} \left\{ \frac{(Mn - mn)^2}{mn(\lambda_{ij}b)} + \frac{1}{mn(\lambda_{ij}b)} \right\} \right) \right] \right\} \end{aligned}$$

where  $H(w) = \int K(w)^2 dw$ ,  $d_k = \int w^2 K(w) dw$ .

From equation (65), it is noted that if the sample size is large, that is as  $n \rightarrow N$  and  $m \rightarrow M$ , the  $MSE$  of  $\hat{\bar{Y}}_{INW}$  due to the kernel tends to zero for a sufficiently small bandwidth. The estimator  $\hat{\bar{Y}}$  is therefore asymptotically consistent since its  $MSE$  converges to zero in probability.

### 3. SIMULATION STUDY

A simulation experiment was conducted using R code in order to compare the performance of the proposed estimator in two-stage cluster sampling with the transformed estimator due to (cf. [5]) and the non-parametric regression estimator due to (cf. [6]). An asymptotic framework is used where both the population number of clusters and the sample number of clusters are large. The number of clusters within each cluster,  $N_i$ , is held constant so that no cluster dominates the population.

Both linear and non-linear mean functions of auxiliary random variables due to (cf. [6]) were considered in generating data, where  $x \in (0, 1)$ . The equations of the mean functions used in simulating the data are given in table 1 below.

TABLE 1. Equations of Mean Functions Simulated

Mean function	Equation
Linear	$m_1(\hat{x}) = 1 + 2(x - 0.5)$
Quadratic	$m_2(\hat{x}) = 1 + 2(x - 0.5)^2$
Sine	$m_3(\hat{x}) = 2 + \sin(2\pi x)$
Exponential	$m_4(\hat{x}) = \exp(-8x)$
Bump	$m_5(\hat{x}) = 1 + 2(x - 0.5) + \exp\{-200(x - 0.5)^2\}$
Jump	$m_6(\hat{x}) = 1 + 2(x - 0.5)I_{x \leq 0.65} + 0.65I_{x \geq 0.65}$

The population auxiliary values,  $x_{ij}$ , of size  $M = 2000$  are generated as identical and independently distributed uniform  $(0, 1)$  random variables. The survey values are only known for the respondents in the selected sample. Using the auxiliary values, the non-response values are generated, that is, for every generated value  $x_{ij}, i = 1, 2, \dots, M; j = 1, 2, \dots, N_i$ , the mean survey non-response values are generated as

$$(66) \quad \hat{y}_{ij} = \frac{1}{M} \left\{ \frac{m(\hat{x}_{ij})}{N_i} + \frac{\hat{e}_{ij}}{N_i} \right\}$$

where  $\hat{e}_{ij}$  are identically and independently distributed normal random variables with mean zero and variance one. Besides, a Gaussian kernel with mean zero and variance one was used. A Gaussian kernel was used since it has smooth and continuous derivatives at every data point. Besides, an optimal bandwidth generated using cross-validation technique due to (cf. [14]) was used. It has been noted by (cf. [14]) that this bandwidth would lead to more informative estimates compared to other choices. The local bandwidth,  $\lambda_{ij}$ , given in equation (??) were generated using the algorithm due to (cf. [12]).

At stage one, a sample of clusters is generated first by simple random sampling using a sample of size  $m = 200$ . At stage two, sub-samples of elements within every selected cluster are generated by simple random sampling with replacement using a random sample of size  $n_i$ . The non-response mean survey values were then generated using equation (66). The estimates of finite population mean were then computed using the estimator in equation (15). The values of bias and mean squared error values were also computed. The 95%

confidence intervals were then constructed for the estimators of the finite population means for comparative purposes.

#### 4. SIMULATION RESULTS

The values of the bias, mean squared error and confidence interval lengths are given in the following tables. Note that  $\hat{\hat{Y}}_{INW}$  is the estimator of finite population mean proposed in this study,  $\hat{\hat{Y}}_{TDM}$  is the transformation of data method estimator of finite population mean due to (*cf.* [5]) whereas  $\hat{\hat{Y}}_{REG}$  is the non-parametric regression estimator due to (*cf.* [6]). Both  $\hat{\hat{Y}}_{TDM}$  and  $\hat{\hat{Y}}_{REG}$  were used for comparative purposes with the proposed estimator.

TABLE 2. Summary Results of Bias

Estimators	$\hat{\hat{Y}}_{INW}$	$\hat{\hat{Y}}_{TDM}$	$\hat{\hat{Y}}_{REG}$
Linear	-0.00213	-0.1667	-0.3312
Quadratic	-0.0132	0.04171	-0.0966
Sine	-0.0521	-0.6416	-1.2311
Exponential	-0.0041	0.3592	0.7225
Bump	-0.0032	-0.2358	-0.4685
Jump	-0.0188	-0.2466	-0.4743

The biases of the estimators considered are presented in table 2 above. Negative values of the bias imply underestimation while positive values of the bias indicate overestimation of the finite population mean by the different estimators. The proposed estimator has relatively smaller values of the bias followed by transformation of data method estimator due to (*cf.* [5]). The non-parametric-based estimator due to (*cf.* [6]) has larger values compared to the other two estimators. It is also observed that the three estimators have relatively closer values of the bias in the quadratic mean function though the transformation of data method has positive bias at this mean function. Generally, among the three estimators of finite population mean, the proposed estimator using improved Nadaraya-Watson kernel regression technique performs better than the other two estimators in terms of bias.

TABLE 3. Summary Results of MSE Values

Estimators	$\hat{\hat{Y}}_{INW}$	$\hat{\hat{Y}}_{TDM}$	$\hat{\hat{Y}}_{REG}$
Linear	0.0334	0.1097	0.1321
Quadratic	0.0093	0.1455	0.5835
Sine	0.4215	1.5157	1.555
Exponential	0.3430	0.5220	1.3780
Bump	0.0634	0.2195	0.2508
Jump	0.2250	0.2951	1.1611

Mean squared error combines both the variance and the squared bias terms of an estimator. The mean squared error values presented in tables 3 were simulated using the different mean functions indicated. The quadratic mean function gives the smallest value of the mean squared error of the proposed estimator followed by the linear function. The estimator due to (*cf.* [6]) has the largest value of the mean squared error in the jump function. Generally, it is noted from table 3 that the mean squared error values for the proposed estimator are relatively smaller than the rest of the estimators considered. The transformation of data method estimator due to (*cf.* [5]) follows closely in the second place with smaller mean squared error values compared to non-parametric regression-based estimator due to (*cf.* [6]). From this comparison of the mean squared error values, it can be concluded that the proposed estimator is more efficient than the other two estimators considered. It has got smaller MSE values in all the mean functions and thus outperforms the others in terms of efficiency.

TABLE 4. Summary Results of 95% CI Lengths

Estimators	$\hat{Y}_{INW}$	$\hat{Y}_{TDM}$	$\hat{Y}_{REG}$
Linear	0.7164	1.2984	1.4249
Quadratic	0.8270	1.4951	2.994
Sine	1.5269	2.5451	4.888
Exponential	2.2958	2.8322	4.6016
Bump	0.9872	1.8365	1.9630
Jump	1.8594	2.1297	4.2239

The 95% upper and lower confidence intervals were constructed for the estimators of finite population mean. Confidence interval lengths were then obtained. The results are given in table 4. From the values obtained, it is noted that the confidence interval lengths for the proposed estimator are much tighter than those of the estimators due to (*cf.* [6]) and (*cf.* [5]). Hence, at 95% level of confidence, the estimator proposed in this study performs better than its rival estimators.

## 5. CONCLUSION

This study has developed an estimator of finite population mean in two-stage cluster sampling assuming random non-response occurs in the survey variable in the second stage of cluster sampling. Complete auxiliary information is assumed to be available in both stage one and stage two of cluster sampling. Kernel weights developed using improved Nadaraya-Watson regression technique were used in the estimation process. The theoretical properties of the proposed estimator such as asymptotic bias, variance and mean squared error were derived. Simulation results show that the proposed estimator has smaller values of the bias, smaller mean squared error values and tighter confidence interval lengths compared to the other estimators. Therefore, the estimator of finite population mean proposed in this study dominates the estimators due to (*cf.* [6]) and (*cf.* [5]) respectively.

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## A note on generation of all Pythagorean triples

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**ABSTRACT.** There exist several techniques used to generate Pythagorean triples. The most effective formula for generating Pythagorean triples is the Euclid's formula. Whereas the Euclid's formula generate infinitely many Pythagorean triples, it does not generate all of them. For instance, the Euclid's formula generates the triple (3, 4, 5) but does not generate (4, 3, 5), in which case a transposition is needed. In addition, the triple (9, 12, 15) cannot be generated directly from the Euclid's formula but rather a multiplier to the triple (3, 4, 5) does so. In this note, we establish a formula which generates all Pythagorean triples, primitive and non-primitive, without using a transformation and without using a multiplier. .

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### 1. INTRODUCTION

A Pythagorean Triple (PT) is a triple of positive integers  $(a, b, c)$ , which satisfies the Pythagorean equation

$$a^2 + b^2 = c^2,$$

where  $c$  represents the length of the hypotenuse,  $a$  and  $b$  represent the lengths of the other two sides (called legs) of a right triangle. In other words a Pythagorean triple represents the lengths of the sides of a right triangle where all the three sides have integer lengths. We say a Pythagorean triple  $(a, b, c)$  is primitive if the numbers  $a, b$  and  $c$  are pairwise co-prime [6].

We first note the parity of  $a, b$  and  $c$  in primitive triples, that is their values modulo 2. Since  $0^2 \equiv 0, 1^2 \equiv 1, 2^2 \equiv 0, \text{ and } 3^2 \equiv 1 \pmod{4}$ , the only squares *modulo* 4 are 0 and 1. Letting  $A = a^2, B = b^2$ , and  $C = c^2$ , we have the following solutions to  $A + B \equiv C \pmod{4}$ :

$$(1) \quad 0 + 0 \equiv 0; 0 + 1 \equiv 1; 1 + 0 \equiv 1.$$

( $1 + 1 \equiv 2$  is not a solution because 2 is not a square modulo 4.) The first of these solutions corresponds only to non-primitive triples where  $a, b$  and  $c$  are all divisible by 2. Therefore any primitive triple corresponds to one of the other two solutions. In either case,  $C$  is odd, and exactly one of  $A$  and  $B$  is always odd. Thus  $c$  is odd, and exactly one of  $a$  and  $b$  is odd.

Many formulas for generating Pythagorean triples with particular properties have been since the time of Euclid and also by several other scholars, see [1, 2, 3, 5, 8, 9, 10, 12]. In [4], integer solutions to the Pythagorean equation are determined by finding positive integers  $r, s$  and  $t$  such that  $r^2 = 2st$ . The proof of this method is given in [14] and it uses a bijection which entails transforming some grid squares into an object that uniquely determines a Pythagorean triple.

In [11], McCullough uses the height and excess enumeration to generate Pythagorean triples. In this method, for a Pythagorean triple  $(a, b, c)$ , the height  $h$  is just  $c - b$ , and the excess  $e$  is  $a + b - c$ . The term excess arises from the fact that  $e$  is simply the extra distance one must travel when going along the two legs instead of the hypotenuse. In [13], Roy and Sonia uses the difference between one leg and hypotenuse to formulate a method of generating Pythagorean triples.



Euclid's formula presents the most common way of generating Pythagorean triples. It states that primitive Pythagorean triples  $(a, b, c)$  in which  $b$  is even, are generated by the formulae

$$a = n^2 - m^2, \quad b = 2nm, \quad c = m^2 + n^2,$$

where  $n > m > 0$ ;  $m, n \in \mathbb{Z}^+$ , for any pair of co-prime positive integers of opposite parity. A triple generated using this method is primitive if  $(n, m) = 1$  and if  $n$  and  $m$  are of opposite parity [6, 15, 16]. In this formula, a convention is made in which  $a$  is always odd and  $b$  is always even, since otherwise we can rename the variables of a given triple to obtain this. Therefore, using this formula, a Primitive Pythagorean Triple has a unique representation  $(a, b, c)$ , where  $b$  is even and  $a$  and  $c$  are odd. To obtain the other set of solutions  $(b, a, c)$ , a transformation is required. We address this problem in the subsequent sections.

## 2. UNIT CIRCLE

In this section, we use a parametrization of the Unit Circle to find rational points on the Unit Circle (point  $(x, y)$  satisfying the equation  $x^2 + y^2 = 1$ , where both  $x$  and  $y$  are rational numbers).

Consider the lines passing through the point  $(-1, 0)$  on the Unit Circle. Let  $t = m$  be the slope of one such a line. Notice that every line through  $(-1, 0)$  intersects the circle at exactly one point. Next, we point out that since the distance from  $(-1, 0)$  to the origin is 1, the line with slope  $t$  will have a  $y$ -intercept of  $(0, t)$ . Thus, except the vertical line  $x = -1$ , the other lines have equations of the form

$$(2) \quad y = tx + t.$$

Moreover, each of these lines cuts the circle in exactly two points, one of which is  $(-1, 0)$ . Let us find the other point.

Substituting (2) into the equation of the Unit Circle gives

$$x^2 + t^2(x + 1)^2 = 1,$$

or

$$(3) \quad (1 + t^2)x^2 + 2t^2x + (t^2 - 1) = 0.$$

Solving (3) for  $x$ , we find

$$\begin{aligned} x &= \frac{-2t^2 \pm \sqrt{(4t^4 - 4(1 + t^2)(t^2 - 1))}}{2(1 + t^2)} \\ &= -1 \text{ or } \frac{1 - t^2}{1 + t^2} \end{aligned}$$

and  $y = t(x + 1) = 0$  or  $t\left(\frac{1 - t^2}{1 + t^2} + 1\right) = \frac{2t}{1 + t^2}$ .

Thus, we have a parametrization of the Unit Circle, excluding the point  $(-1, 0)$ , with parameter  $t$  :

$$(4) \quad x(t) = \frac{1 - t^2}{1 + t^2}; \quad y(t) = \frac{2t}{1 + t^2}.$$

It can easily be checked that this is indeed the Unit Circle by substituting into the Cartesian equation  $x^2 + y^2 = 1$ .

Substituting any rational number for the parameter  $t$  in (4) will give a rational number. For example, if  $t = 1$ , then  $x = 0$  and  $y = 1$ . If  $t = \frac{1}{2}$ , then  $x = \frac{3}{5}$ ,  $y = \frac{4}{5}$  and if  $t = \frac{1}{5}$  then

$x = \frac{24}{26} = \frac{12}{13}$ ,  $y = \frac{10}{26} = \frac{5}{13}$ , and so on. Each rational point yields a Pythagorean Triple. For example,

$$\left(\frac{3}{5}\right)^2 + \left(\frac{4}{5}\right)^2 = 1$$

implies that (3, 4, 5) is a Pythagorean Triple;

$$\left(\frac{24}{26}\right)^2 + \left(\frac{10}{26}\right)^2 = 1$$

implies that (24, 10, 26) is a Pythagorean Triple; and the simplified form of this equation

$$\left(\frac{12}{13}\right)^2 + \left(\frac{5}{13}\right)^2 = 1$$

yields the Pythagorean Triple (12, 5, 13).

PYTHAGOREAN TRIPLES

Let  $(x, y) = \left(\frac{1-t^2}{1+t^2}, \frac{2t}{1+t^2}\right)$  be a parametrization of the Unit Circle  $x^2 + y^2 = 1$ , where  $t$  is the slope of any line through the point  $(-1, 0)$  and cutting the circle at  $(x, y)$ . For any  $t$  with  $0 < t < 1$ , the parametric equations  $x(t)$  and  $y(t)$  yield Pythagorean Triples  $(a, b, c)$ . In the table below, we enumerate some Triples for  $t = \frac{r}{s} < 1$ , where  $0 < r < s < 10$ , and  $(r, s) = 1$ .

$s$	$r$	$t$	$x$	$y$	$PT = (a, b, c)$	$PPT = (a, b, c)$
2	1	$\frac{1}{2}$	$\frac{3}{5}$	$\frac{4}{5}$	(3, 4, 5)	(3, 4, 5)
3	1	$\frac{1}{3}$	$\frac{8}{10}$	$\frac{6}{10}$	(8, 6, 10)	(4, 3, 5)
3	2	$\frac{2}{3}$	$\frac{5}{13}$	$\frac{12}{13}$	(5, 12, 13)	(5, 12, 13)
4	1	$\frac{1}{4}$	$\frac{15}{17}$	$\frac{8}{17}$	(15, 8, 17)	(15, 8, 17)
4	3	$\frac{3}{4}$	$\frac{7}{25}$	$\frac{24}{25}$	(7, 24, 25)	(7, 24, 25)
5	1	$\frac{1}{5}$	$\frac{24}{26}$	$\frac{10}{26}$	(24, 10, 26)	(12, 5, 13)
5	2	$\frac{2}{5}$	$\frac{21}{29}$	$\frac{20}{29}$	(21, 20, 29)	(21, 20, 29)
5	3	$\frac{3}{5}$	$\frac{16}{34}$	$\frac{30}{34}$	(16, 30, 34)	(8, 15, 17)
6	1	$\frac{1}{6}$	$\frac{35}{37}$	$\frac{12}{37}$	(35, 12, 37)	(35, 12, 37)
6	5	$\frac{5}{6}$	$\frac{11}{61}$	$\frac{60}{61}$	(11, 60, 61)	(11, 60, 61)
7	1	$\frac{1}{7}$	$\frac{48}{50}$	$\frac{14}{50}$	(48, 14, 50)	(24, 7, 25)
7	2	$\frac{2}{7}$	$\frac{45}{53}$	$\frac{28}{53}$	(45, 28, 53)	(45, 28, 53)
7	3	$\frac{3}{7}$	$\frac{40}{58}$	$\frac{42}{58}$	(40, 42, 58)	(20, 21, 29)
7	4	$\frac{4}{7}$	$\frac{33}{65}$	$\frac{56}{65}$	(33, 56, 65)	(33, 56, 65)
8	1	$\frac{1}{8}$	$\frac{63}{65}$	$\frac{16}{65}$	(63, 16, 65)	(63, 16, 65)
8	3	$\frac{3}{8}$	$\frac{55}{73}$	$\frac{48}{73}$	(55, 48, 73)	(55, 48, 73)
8	5	$\frac{5}{8}$	$\frac{39}{89}$	$\frac{80}{89}$	(39, 80, 89)	(39, 80, 89)
8	7	$\frac{7}{8}$	$\frac{15}{113}$	$\frac{112}{113}$	(15, 112, 113)	(15, 112, 113)
9	1	$\frac{1}{9}$	$\frac{80}{82}$	$\frac{18}{82}$	(80, 18, 82)	(40, 9, 41)
9	2	$\frac{2}{9}$	$\frac{77}{85}$	$\frac{36}{85}$	(77, 36, 85)	(77, 36, 85)
9	4	$\frac{4}{9}$	$\frac{65}{97}$	$\frac{72}{97}$	(65, 72, 97)	(65, 72, 97)
9	5	$\frac{5}{9}$	$\frac{56}{106}$	$\frac{90}{106}$	(56, 90, 106)	(28, 45, 53)
9	7	$\frac{7}{9}$	$\frac{32}{130}$	$\frac{126}{130}$	(32, 126, 130)	(16, 63, 65)
9	8	$\frac{8}{9}$	$\frac{17}{145}$	$\frac{144}{145}$	(17, 144, 145)	(17, 144, 145)
10	1	$\frac{1}{10}$	$\frac{99}{101}$	$\frac{20}{101}$	(99, 20, 101)	(99, 20, 101)
10	3	$\frac{3}{10}$	$\frac{91}{109}$	$\frac{60}{109}$	(91, 60, 109)	(91, 60, 109)
10	7	$\frac{7}{10}$	$\frac{51}{149}$	$\frac{140}{149}$	(51, 140, 149)	(51, 140, 149)
10	9	$\frac{9}{10}$	$\frac{19}{181}$	$\frac{180}{181}$	(19, 180, 181)	(19, 180, 181)

Table 1. Triples for values of  $t$  with  $0 < r < s < 10$ .

**Note 2.1.** We note that this method is incapable of generating all the Pythagorean triples. For instance, the triple (4, 3, 5) cannot be generated directly. A transposition on the triple (3, 4, 5) is applied to obtain (4, 3, 5). In addition the triple (9, 12, 15) requires a multiplier on (3, 4, 5). We establish a formula that generates all Pythagorean triples without need for either transpositions or multipliers.

**Theorem 2.2.** Let  $(a, b, c)$  be a primitive Pythagorean triple with  $b$  even. Then there exist relatively prime positive integers  $m < n$  having distinct parities (i.e. one even and one odd) such that

$$(5) \quad a = n^2 - m^2, \quad b = 2nm, \quad c = n^2 + m^2.$$

In Table 2, we generate a few examples of Pythagorean triples using equation (5). We can obtain an equivalent formula to Theorem 2.2 from the equation  $a^2 + b^2 = c^2$  by writing

$$a^2 = c^2 - b^2 = (c - b)(c + b)$$

in place of  $b^2 = c^2 - a^2$ , for a given primitive Pythagorean triple  $(a, b, c)$ .

$n$	$m$	$a = n^2 - m^2$	$b = 2nm$	$c = n^2 + m^2$	$DL(c, a) = 2m^2$	$DL(c, b) = u^2$
2	1	3	4	5	2	1
3	1	8	6	10	2	4
3	2	5	12	13	8	1
4	1	15	8	17	2	9
4	2	12	16	20	8	4
4	3	7	24	25	18	1
5	1	24	10	26	2	16
5	2	21	20	29	8	9
5	3	16	30	34	18	4
5	4	9	40	41	32	1

**Table 2.** Examples of Pythagorean triples generated from equation (5).

**Theorem 2.3.** Every primitive Pythagorean triple  $(a, b, c)$  with  $a$  odd and  $b$  even can be obtained by using the formulas

$$a = st, \quad b = \frac{s^2 - t^2}{2}, \quad c = \frac{s^2 + t^2}{2},$$

where  $s > t \geq 1$  are chosen to be any odd integers with no common factors.

Notice that Theorems 2.2 and 2.3 are equivalent under the substitution  $t = n - m$ ,  $s = n + m$ .

**Remark 1.** Let

$$(a, b, c) = (n^2 - m^2, 2nm, n^2 + m^2),$$

be a primitive Pythagorean triple, where  $n > m$ ,  $(n, m) = 1$  and  $n$  and  $m$  are of opposite parity.

The two methods in Theorems 2.2 and 2.3, like many others will require a transposition. For instance in both  $b$  is always even whenever  $(a, b, c)$  is primitive.

In the following section, we consider a formula which is able to generate all Pythagorean triples without need for either multipliers or transpositions.

3. GENERATING ALL PYTHAGOREAN TRIPLES

It can be seen that the difference formula does not generate all Pythagorean triples. For instance the triple (9, 12, 15) cannot be generated directly by the formula but rather applying a multiplier of 3 to the triple (3, 4, 5) does so. Moreover, for every primitive triple  $(a, b, c)$  generated by the formula, the corresponding primitive triple  $(b, a, c)$  cannot be directly generated, a transposition is required.

If  $(a, b, c)$  is a Pythagorean triple, the triple  $(b, a, c)$  is called its co-triple. In this section, we illustrate an expression which generates all Pythagorean triples uniquely, as discussed below.

**Proposition 3.1.** *Let  $(a, b, c)$  be a Pythagorean triple. Then every Pythagorean triple can be generated by the formula:*

$$(6) \quad (a, b, c) = \left( \sqrt{PQ}, \frac{P-Q}{2}, \frac{P+Q}{2} \right),$$

for some positive integers  $P, Q$ .

*Proof.* A Pythagorean triple  $(a, b, c)$  satisfies the equation  $a^2 + b^2 = c^2$ . From this we see that  $a^2 = c^2 - b^2 = (c+b)(c-b)$ . Let  $P = c+b$  and  $Q = c-b$ . Clearly  $P > Q$ . Solving for  $c$  and  $b$ , one obtains:

$$P + Q = 2c \quad \Leftrightarrow \quad c = \frac{P+Q}{2}$$

and

$$P - Q = 2b \quad \Leftrightarrow \quad b = \frac{P-Q}{2}.$$

Since  $P$  and  $Q$  are integers,  $P$  and  $Q$  are either both even or both odd. So for each  $a$ ,  $a^2$  is a product of two factors  $P$  and  $Q$ , with  $P > Q > 0$ .

Next, suppose  $c = b + d$  where  $d$  is the difference between  $c$  and  $b$ , then

$$a^2 + b^2 = (b+d)^2,$$

which simplifies to

$$a^2 = 2bd + d^2$$

or

$$(7) \quad a^2 = d(2b + d).$$

Substitute  $\frac{P-Q}{2}$  for  $b$  in (7) to obtain:

$$\begin{aligned} a^2 &= d[2(P-Q) + d] = PQ \\ \Leftrightarrow dP - dQ + d^2 &= PQ \\ \Leftrightarrow dP + d^2 &= (P+d)Q \\ \Leftrightarrow d(P+d) &= (P+d)Q. \end{aligned}$$

Thus  $d = Q$ . So  $a = \sqrt{PQ}$ , where  $P > Q$  and positive integers  $P, Q$  are either both odd or both even. Then

$$(a, b, c) = \left( \sqrt{PQ}, \frac{P-Q}{2}, \frac{P+Q}{2} \right).$$

Next, we show that this is a Pythagorean triple, that is,

$$\begin{aligned} \left( \sqrt{PQ} \right)^2 + \left( \frac{P-Q}{2} \right)^2 &= PQ + \frac{P^2 - 2PQ + Q^2}{4} \\ &= \frac{P^2 + 2PQ + Q^2}{4} \\ &= \left( \frac{P+Q}{2} \right)^2. \end{aligned}$$

□

For every primitive Pythagorean triple  $(a, b, c)$  one can easily obtain the corresponding co-triple  $(b, a, c)$  from (6) as illustrated in the corollary below:

**Corollary 3.2.** *If  $P = u^2$  and  $Q = v^2$  where  $u, v$  are relatively prime odd positive integers, such that  $u > v$ , then  $(a, b, c)$  is a primitive Pythagorean triple with  $a$  odd. The corresponding co-triple is generated when  $P = 2 \left(\frac{u+v}{2}\right)^2$  and  $Q = 2 \left(\frac{u-v}{2}\right)^2$ .*

*Proof.* Let  $P = u^2$  and  $Q = v^2$ , then by Proposition 3.1, we obtain

$$(8) \quad (a, b, c) = \left( uv, \frac{u^2 - v^2}{2}, \frac{u^2 + v^2}{2} \right).$$

Clearly, if  $u > v > 0$  are odd integers such that  $(u, v) = 1$ , then (8) is a primitive Pythagorean triple with  $a$  odd.

If  $P = 2 \left(\frac{u+v}{2}\right)^2$  and  $Q = 2 \left(\frac{u-v}{2}\right)^2$ , then

$$\sqrt{PQ} = \sqrt{2 \left(\frac{u+v}{2}\right)^2 \cdot 2 \left(\frac{u-v}{2}\right)^2} = \frac{u^2 - v^2}{2} = b,$$

$$\frac{P - Q}{2} = \sqrt{\frac{2 \left(\frac{u+v}{2}\right)^2 - 2 \left(\frac{u-v}{2}\right)^2}{2}} = \frac{u^2 + 2uv + v^2 - u^2 + 2uv - v^2}{4} = \frac{4uv}{4} = uv = a,$$

and

$$\frac{P + Q}{2} = \sqrt{\frac{2 \left(\frac{u+v}{2}\right)^2 + 2 \left(\frac{u-v}{2}\right)^2}{2}} = \frac{u^2 + 2uv + v^2 + u^2 - 2uv + v^2}{4} = \frac{u^2 + v^2}{2} = c.$$

□

Non-primitive Pythagorean triples can be generated as described below:

**Corollary 3.3.** *If  $P$  and  $Q$  are not relatively prime, then  $(a, b, c)$  corresponding  $(b, a, c)$  are non-primitive Pythagorean triples.*

*Proof.* Let  $P = ku^2$  and  $Q = kv^2$ , then by (6),

$$\left( \sqrt{ku^2 \cdot kv^2}, \frac{ku^2 - kv^2}{2}, \frac{ku^2 + kv^2}{2} \right) = k \cdot \left( uv, \frac{u^2 - v^2}{2}, \frac{u^2 + v^2}{2} \right) = k(a, b, c).$$

Similarly, if  $P = 2k \left[\frac{u+v}{2}\right]^2$  and  $Q = 2k \left[\frac{u-v}{2}\right]^2$ , then

$$\begin{aligned} & \left( \sqrt{2k \left[\frac{u+v}{2}\right]^2 \cdot 2k \left[\frac{u-v}{2}\right]^2}, \frac{2k \left[\frac{u+v}{2}\right]^2 - 2k \left[\frac{u-v}{2}\right]^2}{2}, \frac{2k \left[\frac{u+v}{2}\right]^2 + 2k \left[\frac{u-v}{2}\right]^2}{2} \right) \\ &= \left( 2k \left[\frac{u+v}{2}\right] \cdot \left[\frac{u-v}{2}\right], 2k \cdot \frac{4uv}{8}, 2k \cdot \frac{2u^2 + 2v^2}{8} \right) \\ &= k \cdot \left( \frac{u^2 - v^2}{2}, uv, \frac{u^2 + v^2}{2} \right) \\ &= k(b, a, c). \end{aligned}$$

□

Tables 3 and 4 below shows examples of primitive and non-primitive triples and their co-triples generated by this method:

$u$	$v$	$P$	$Q$	$a$	$b$	$c$	$k$	$m$	$P$	$Q$	$b$	$a$	$c$
3	1	9	1	3	4	5	2	1	8	2	4	3	5
5	1	25	1	5	12	13	3	2	18	8	12	5	13
7	1	49	1	7	24	25	4	3	32	18	24	7	25
9	1	81	1	9	40	41	5	3	50	32	40	9	41
5	3	25	9	15	8	17	4	1	32	2	8	15	17
7	3	49	9	21	20	29	5	2	50	8	20	21	29
11	3	121	9	33	56	65	7	4	98	72	56	33	65
13	3	169	9	39	80	89	8	5	128	50	80	39	89

**Table 3.** Primitive Pythagorean Triples and their co-triples

$P$	$Q$	$a$	$b$	$c$	$P$	$Q$	$b$	$a$	$c$
18	2	6	8	10	16	4	8	6	10
27	3	9	12	15	24	6	12	9	15
36	4	12	16	20	32	8	16	12	20
45	5	15	20	25	40	10	20	15	25
50	18	30	16	34	64	4	16	30	34
98	18	42	40	58	100	16	40	42	58
242	18	66	112	130	196	144	112	66	130
338	18	78	160	178	256	100	160	78	178

**Table 4.** Non-primitive Pythagorean Triples and their co-triples

#### 4. APPLICATIONS IN CRYPTOGRAPHY

**Theorem 4.1.** Primitive Pythagorean triples come in 6 classes based on the divisibility of  $a$ ,  $b$ ,  $c$  by 3, 4, and 5.

- Proof.*
- 1 Class A:  $a$  is divisible by 3 and  $c$  is divisible by 5. e.g (3,4,5), (33,56,65)
  - 2 Class B:  $a$  is divisible by 5, and  $b$  is divisible by 3. e.g (5,12,13)
  - 3 Class C:  $a$  is divisible by 3 and 5. e.g (15,8,17), (45,28,53)
  - 4 Class D:  $b$  is divisible by 3 and  $c$  by 5. e.g (7,24,25), (13,84,85)
  - 5 Class E:  $a$  is divisible by 3 and  $b$  by 5. e.g (21,20,29), (9,40,41)
  - 6 Class F:  $b$  is divisible by 3 and 5. e.g (11,60,61), (91,60,109)

□

An experiment was done where 4448 primitive Pythagorean triples were generated by Euclid's formula, which is limited in the Pythagorean Triples it generates, and indexed by increasing  $a$ ,  $b$ , and  $c$  respectively.

Indexed by increasing  $a$ :

ABDEFDCCDFEEDBAFFAABBDEEFDCDDFEEDBBAFFFAABBDEEFDCDDFEDBB  
AAFFFAABBDEEFDCDDCFEEDBBAAFFFAABDEEFDDCCDFEEDDBBAA...

Indexed by increasing  $b$ :

ACBBAEEDDCCDDEEAABBCCAAFFFAACBBAAEEDDDCCCDDEEAABBCCAA  
FFFAACBBBBAAEEDDCCCDDEEAABBBBCCAAFFFAACBBAAEEDDD...

Indexed by increasing  $c$ :

ABCDEBECFAABDDEBEFCACDDEBFCFAABCDDEEFDFCDDEBEFCFAABCDDBFC  
AABCCEBEFFAABDDEBEFFAABDDDEECCFAACDDEBFCFBCDEEECFBCDDEEBB  
EFAABDDBEFCA...

We obtain separate sequences related to the occurrence of As, Bs, Cs, Ds, Es, and Fs by considering the distance between occurrences of the letters. Thus in the listing by increasing  $c$ , A occurs, after its first value, at the 10th, 11th, 20th, ... positions, which corresponds to



the numbers 9, 1, 9, . . . . These sequences are called Baudhāyana sequences.

However, equation (6) makes it easy to generate and conveniently classify all triples into these classes. This helps improve the randomness property of Baudhāyana sequences which have applications in key distribution and in information hiding in the field of Cryptography.

For instance, consider the triples generated by (6) and indexed by increasing  $a$ . We obtain the following

$a$	$b$	$c$	Class
3	4	5	A
4	3	5	D
5	12	13	B
7	24	25	D
8	15	17	F
9	40	41	E
11	60	61	F
12	5	13	E
12	35	37	E
13	84	85	D
15	8	17	C
16	63	65	D
17	144	145	D
19	180	181	F
20	21	29	B
20	99	101	B
21	20	99	E
21	220	221	E
23	264	265	D
24	7	25	A
24	143	145	A

**Table 5.** Primitive Pythagorean Triples by increasing values of  $a$

Thus in the listing by increasing  $a$ , A occurs, after its first value, at the 20th, 21st which correspond to the numbers 1, 18, 1, . . . .

Similarly in the listing by increasing  $c$ , A occurs, after its first value, at the 8th, 18th which correspond to the numbers 1, 7, 17, . . . .

#### CONCLUSION

Proposition 3.1 presents an effective way of generating all Pythagorean triples without need for multipliers or transpositions. It is easy to determine the values of the generating pair of integers  $P$  and  $Q$  such that (6) is a triple. If  $u > v$  are relatively prime odd positive integers such that  $u^2 = P$  and  $v^2 = Q$  then  $P, Q$  are odd and  $(P, Q) = 1$ . In this case  $a$  is odd and  $(a, b, c)$  is primitive. Similarly, if  $k > m$  are any relatively prime positive integers such that  $P = 2k^2$  and  $Q = 2m^2$ , then  $(P, Q) = 2$ . Clearly  $a$  is even and  $(a, b, c)$  is primitive. To obtain non-primitive triples, one may use the expression  $n(P, Q)$  for  $n \in \mathbb{Z}^+$ .

This formula makes it easy to generate and conveniently classify triples into classes according to divisibility by 3 or 5 or both. This helps improve the randomness property and its application to Cryptography.

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